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How Partisanship in Cities Influences Housing Policy

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Abstract

Housing policy is one of the most important areas of local politics. Yet little is known about how local legislatures and executives make housing policy decisions and how their elections shape policy in this important realm. We leverage housing policy data and a new data source of 15,520 city council elections and 3,238 mayoral elections in large cities in the United States and a regression discontinuity design to examine partisan divides in housing policy among the mass public as well as the impact of local leaders' partisanship on housing policy. We provide robust evidence that electing mayors from different political parties shapes cities' housing stock. Electing a Democrat as mayor leads to increased multifamily housing production. These effects are concentrated in cities where councils have less power over land use changes. Overall, our paper shows that politics influences local housing policy, and it contributes to a larger literature on local political economy.

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Political opposition to new higher density housing and restrictive local land use policy have combined to block residential housing construction in many metropolitan areas where it is most needed (Einstein, 2021; Einstein, Glick, and Palmer, 2019; Trounstine, 2020). This has important consequences. Across a number of metrics, a housing affordability crisis has not only reached higher peaks in recent years, but has also spread to parts of the country previously less affected – no longer impacting only very large and coastal cities (Corinth and Dante, 2022; Kingsella and MacArthur, 2022). Median home price-to-income ratios in the early 2020's are 30% higher than they were in the 2000's. Both the supply of new homes on the market and the stock of lower-rent rental units are at all time lows.¹ Substantial evidence exists that restrictions on supply contribute to higher housing prices (Gyourko and Molloy, 2015). Beyond direct impacts on affordability, an undersupply of affordable housing may negatively impact local economies as well as more generally exacerbate economic inequality within and across cities (Hsieh and Moretti, 2019; Saks, 2008).

As a result of the housing crisis, some of the most contentious policymaking debates in large cities are now focused on residential housing development. Because these large cities tend to be more liberal and Democratic, many in the popular media have suggested that Democrats — both elites and members of the mass public — have caused this housing crisis (e.g. McArdle, 2018). On the other hand, there is evidence from surveys of both the mass public and city leaders that Democrats are more likely than Republicans to support dense, multifamily housing (e.g., Jones, 2020; Einstein, Glick, and Palmer, 2017; Einstein et al., 2018). This suggests that the partisanship of elected officials in city governments should affect housing policy, and the election of Democrats should lead to more multifamily housing.

While it may seem obvious that partisanship should affect local housing policy, as it does most other aspects of American politics, a variety of previous literature has argued that housing policy differs from more traditionally-ideological fiscal policy. A number of scholars

¹These figures drawn from the Harvard Joint Center for Housing Studies' 2022 reports on the State of the Nation's Housing and Rental Markets

have argued that while partian polarization may exist on abstract housing policies, such divisions might not exist for concrete housing policy decisions and outcomes (Anzia, 2021; Hankinson, 2018). Others have argued that divisions – both at the elite level and among the mass public – on local policy issues like housing may correspond not with partianship but with other factors such as homeownership (Einstein, Glick, and Palmer, 2019), selfinterest (de Benedictis-Kessner and Hankinson, 2019; Marble and Nall, 2021), race and class (Hajnal and Trounstine, 2014; Schaffner, Rhodes, and La Raja, 2020), preferences for privatization (Bucchianeri et al., 2021), or membership in other groups (e.g. Anzia, 2011). Thus the election of politicians from different parties might lead to little changes in the policy preferences of officials in power or policy outcomes. Finally, there is a large literature on the constraints on local officials' influence over policy due to state control, which may limit the ability of local politicians to change policy even if they wished to do so (Peterson, 1981, 1995). Alongside state-imposed constraints, there are also numerous layers of local approval needed for increases in housing development – ranging from appointed planning and zoning commissions, to boards of zoning appeals, to judicial review (Anderson and Sass, 2004). Together, this work suggests that the partial political leaders may not influence housing policy.

In this paper, we leverage a bevy of data sources to test these competing ideas and holistically assess the degree to which the election of city leaders from different parties delivers different housing policy to constituents. Using a new data source of elections in large cities in the United States and a regression discontinuity design, we examine the causal impact of city councilors' and mayors' partisanship on housing policy. We examine 15,520 individual city council elections and 3,238 mayoral elections in 381 large cities between 1990 and 2022. Reflecting previous work on the limits to the role of partisanship in local politics, we provide evidence that electing city councilors from different political parties has no detectable effect on cities' housing policy outcomes. However, we provide robust evidence that mayoral partisanship has large and significant effects on housing outcomes. Electing a Democrat as mayor leads to increases in the supply of multifamily housing units. These effects are larger in places where the city council has no veto power over new development. Moreover, we find suggestive evidence that mayoral partisanship influences the affordability of cities: electing a Democrat as mayor appears to lower housing prices.

Overall, our results indicate that partisan selection has an important influence on city housing policy. Previous work has indicated the *fiscal* policy consequences of local officials' partisanship (de Benedictis-Kessner and Warshaw, 2016, 2020). But it has been unclear whether partisanship in local governments extends to other areas of policy. Here, we show that mayoral partisanship affects contentious multifamily housing production. Thus our results reinforce longstanding wisdom in the urban politics literature on the power of mayors relative to other local officials (e.g. Svara, 1990). Together, our results examining both city councilors and mayors help to provide an updated and holistic assessment of the consequences of partisanship in cities for a policy area of contentious contemporary debate – and contribute to a larger literature on democratic representation in local governments (Trounstine, 2010; Warshaw, 2019).

Theoretical Framework

In this section, we discuss our theoretical framework. We begin by examining the role of partisanship in local governments. We then examine partisan preferences on housing policy. We show that Democrats are more likely to favor dense, multifamily housing than Republicans. We then show that mayors and, to a lesser degree, city councils have the power to influence housing policy in cities. This leads to the expectation that electing Democratic officials will lead to the development of more multifamily housing.

Partisanship in Local Governments

An older literature on urban politics argued that urban policy issues such as paving streets and repairing street lights are inherently apartisan. But a variety of recent studies have found that partisan cleavages in local politics increasingly echo partisan splits in national politics (Hopkins, 2018; Martin and McCrain, 2019). Einstein and Glick (2018) show that Democratic mayors tend to have more liberal preferences on fiscal issues. Bucchianeri et al. (2021) find that partisanship is one of the key factors structuring the ideological preferences of local elected officials. And Lee, Landgrave, and Bansak (2022) find that there are large partisan divisions among local elected officials even after controlling for the voting behavior of their constituents in presidential elections.

Electing individual politicians from one party or another therefore may have the potential to both change the ideology of the mayor as well as the ideological position of the median voter in local legislatures. This could be especially true in policy areas where the ideological preferences of Democratic legislators are far from the preferences of Republican legislators (Lee, Landgrave, and Bansak, 2022). Several recent studies have found that the partisanship of mayors and county legislators affects local policy outputs (de Benedictis-Kessner and Warshaw, 2016, 2020). The influence of partisanship on policy outcomes might naturally extend to city councils as well. Previous research has found that legislators tend to form ideological coalitions even within nominally nonpartisan city councils (Burnett, 2019).²

Partisan Preferences on Housing Policy

Like many other issues, housing policy, and especially multifamily housing production, has increasingly become associated with the two major political parties in the United States, especially as a response to housing supply shortages and the associated impact on local inequality. The Pew Research Center finds that "Republicans and Democrats express sharply different preferences about their ideal communities and house sizes" (Jones, 2020). In a 2019

 $^{^{2}}$ Though, see Bucchianeri (2020).

survey, they found that 58% of Democrats preferred to live in a community where the houses are "smaller and closer to each other" rather than "larger and farther apart." In contrast, only 34% of Republicans share this view.³ This partisan polarization on housing preferences is similar in magnitude in previous versions of this Pew survey from 2014 and 2017. Marble and Nall (2021) also find large partisan gaps in Americans' housing policy preferences. They find that roughly 70% of Democrats believe that the Federal government should ensure all Americans have housing, compared to 33% of Republicans. Finally, Einstein, Palmer, and Glick (2019) find that Democrats are more likely than Republicans to support new housing in local zoning meetings in Massachusetts.⁴

The partisan splits on policy concerning multifamily housing development, in particular, are further supported by an original survey we ran measuring policy preferences at the individual level.⁵ In Figure 1 we show the percentage of respondents, broken down by their partisanship, who supported more multifamily housing development everywhere in their city or town. We find that Democratic respondents are more supportive of multifamily housing development than Republican respondents by 19 percentage points.⁶

There are a variety of theoretical factors that could lead Democrats to be more likely to favor multifamily housing than Republicans. Democratic voters are more likely than Republicans to be renters rather than homeowners (Yoder, 2020; Marble and Nall, 2021). They are also more likely to live in dense areas. According to Brown and Enos (2021, p.

³One concern is that Democrats are more likely to live in cities, and these results could simply reflect where people live. However, when we reanalyzed the Pew survey data, we continued to find a large partial split in multivariate regression models that control for whether people live in a rural areas.

⁴There is also some evidence, especially from studies of the 1990s and early 2000s, that Republicans are more likely than Democrats to support single family housing development in California (Gerber and Phillips, 2003; Kahn, 2011). Gerber and Phillips (2003) argues this reflects Democrats' environmental preferences in favor of retaining open space.

⁵We ran an original survey in November 2021 on a diverse, national sample of US respondents recruited via the firm PureSpectrum, which pays respondents from a marketplace of other survey sample panel companies. We fielded this survey of 2,212 respondents using quotas to match the demographics of the United States as a whole. Among other questions, we asked respondents a question on the development of multifamily housing. Specifically, we asked whether respondents support allowing multifamily housing everywhere in their city or town rather than in specific downtown neighborhoods to gauge public opinion on this contentious policy issue. The full wording of questions used in the survey are in Appendix A.

⁶A two-sample t-test for the difference in proportions indicates that this difference is statistically significant, p = 0.



Figure 1: Partisan Differences in Housing Policy Opinions

1003), 61% of Democrats live in medium or high density Census tracts compared to 39% of Republicans. This could make Democratic voters more sympathetic to multifamily housing. Mass attitudes could also be influenced by the fact that, due in part to decades of racial segregation in real estate, racial minorities are more likely to live in multifamily housing (Anacker, 2018; Trounstine, 2018, 2020). This could make more racially conservative voters, who tend to be Republicans, less likely to support multifamily housing (see Trounstine, 2020). Along similar lines, Pew Research argues that one explanation for partisan splits on housing density is that "partisans diverge on whether it is important that a community is racially and ethnically diverse" (Jones, 2020). Finally, Democrats could be more likely to support multifamily housing due to its environmental benefits over single-family housing (Berrill, Gillingham, and Hertwich, 2021).

The partisan splits on housing policy are also reflected among municipal elected officials. According to a recent report from the *Menino Surveys of Mayors* conducted by Boston University's Initiative on Cities, Democratic and Republican mayors in large cities have different views on housing policy (Einstein, Glick, and Palmer, 2017; Einstein et al., 2018). The 2018 Menino Survey, for example, reports that 68 percent of Democratic mayors supported increasing housing density in established neighborhoods compared to 32 percent of Republicans. In a similar vein, the 2017 Menino Survey indicated that Democratic mayors were more likely to favor "housing stability for renters" and the development of "affordable multi-bedroom units," while Republicans were slightly more likely to favor increasing home ownership rates. One reason that Democratic elected officials are more likely to support multifamily housing could be that re-election oriented officials are aware that people in multifamily housing are more likely to be Democrats (Brown and Enos, 2021). Thus, building more multifamily housing could lead to an increase in the number of Democratic voters, thereby improving Democratic candidates' future electoral prospects.⁷

Finally, there is evidence from the behavior of Democrats and Republicans at the state and national levels of partisan splits on housing. At the national-level, former President Donald Trump argued in 2020 that Democrats would "destroy" suburbs and criticized Democratic attempts to increase multifamily housing.⁸ And in California, Democratic state legislators were much more likely than Republicans to support a recent landmark bill to expand housing production.⁹

The Influence of Local Politicians on Housing Policy

Of course, these partian splits would not necessarily matter if local elected officials had no levers to influence housing policy. But there are a variety of ways that local elected officials are able to influence housing policy. Mayors often set policy goals more broadly for cities using their budget and political leadership. In addition, alongside their local housing authorities, mayors often control the disbursement of federal housing subsidies to local residents and property developers. Mayors also ordinarily appoint members of zoning boards of appeal that control variances to existing zoning regulations.

City councilors may also play an especially important role in housing development given that zoning boards must sometimes seek approval for their decisions from the city council (Anderson, Brees, and Reninger, 2008). City councilors may have incentives to influence development specifically in their geographic districts due to concentrated constituent pressure

⁷Along similar lines, Schmutz and Verdugo (2021) show that left-leaning local elected officials strategically increase the number of left-leaning voters, largely by building more public housing.

⁸For examples of these claims, see Trump, Donald J., and Ben Carson. 2020. "We'll Protect America's Suburbs." Wall Street Journal (August 16), online: https://www.wsj.com/articles/well-protect-americas-suburbs-11597608133, and Keith, Tamara. 2020. "Down in the Polls, Trump Pitches Fear: 'They Want To Destroy Our Suburbs.'" NPR (July 22), online: https://www.npr.org/2020/07/22/893899254/down-in-the-polls-trump-pitches-fear-they-want-to-destroy-our-suburbs.
⁹See https://legiscan.com/CA/votes/SB9/2021.

(Hankinson and Magazinnik, 2022; Mast, 2020).¹⁰ We therefore anticipate that the election to city councils and mayors' offices of a Democrat, rather than Republican, will increase new housing production, especially for multifamily units.

Local officials' ability to influence housing policy is shaped by the local regulatory environment, which varies substantially across the country (Gyourko, Saiz, and Summers, 2008; Gyourko and Krimmel, 2021). The regulatory environment governing housing policy is multidimensional. Zoning plans dictate the types of structures (number of units, height, etc.) that can be built in various parts of a city. But, more pertinent to our study, cities also vary in which local officials are involved in approving new development. In most cities, and certainly in the higher-population cities in our study, city council, commission, or manager approval is required to approve new development in an area where a zoning variance is required – e.g., multifamily housing in a single-family zoned area or a six-story building in an area that only allows three-stories. But, per the Wharton Index, roughly half of cities *also* require approval from one or more of these entities for new "by right" development; that is, development that conforms to local zoning regulations. This type of rule contributes to the stringent regulatory regimes that limit the permitting of multifamily housing and increase the price of land and housing in cities (Kok, Monkkonen, and Quigley, 2014; Gyourko and Krimmel, 2021).

We expect that mayors will be less able to influence housing policy in cities with this added layer of review for by-right development. The effect of this type of rule on the influence of city councilors, on the other hand, is less clear. If the city council's approval power is the singular hindrance to development, an additional member in favor of development would result in more approval under this rule. However, even in cities without this type of council veto power, councils are only one of several bodies with the power to block approval of byright projects; as such their influence may be limited even in places without stringent council majority approval requirements.

¹⁰Though see Gabbe and Kahn (2021) for evidence that this kind of political influence does not happen among Los Angeles city council districts.

Data and Research Design

To assess the effects of partial poince of local housing policy, we leverage a variety of sources of original data. We collect data on city mayoral and legislative elections and housing policy based on permits for private housing development.

City Election Data

To examine the effect of partisan control of local governments on housing policy, we gather data on elections from 1990-2022 in medium and large cities with a population of more than 75,000 people in 2020. We focus on medium and large cities because these cities are likely to have more flexibility to change policy than smaller ones, in part because multifamily housing is more prevalent in medium and larger cities.¹¹ Our dataset consists of 15,520 individual city council elections and 3,238 mayoral elections in 396 cities from 1990 to 2022. The cities in which we have elections data encompass 99% of the population in our target universe of cities over 75,000 in population.¹²

While many big cities have predominantly Democratic populations (Rodden, 2018), their representatives on city councils and in mayors' offices are not uniformly Democratic. In Figure 2 we plot the proportion of council seats held by each party (on the left) and the proportion of mayors from each party (on the right) in the cities in our data over time. While Democrats consistently hold a larger proportion of seats on city councils than Republicans on average, Republicans do hold a nonzero share of the seats on city councils – and close to a quarter of seats on average in the last ten years. Meanwhile, only slightly more than half of the mayors' offices in our data are occupied by Democrats, while around a quarter are occupied by Republicans.

¹¹We use a combination of data from previous papers (Ferreira and Gyourko, 2009; Gerber and Hopkins, 2011; de Benedictis-Kessner and Warshaw, 2016), administrative records, and the crowd-sourced OurCampaigns website to assemble election returns. For officially nonpartisan elections, we estimate candidates' partisanship based on information in OurCampaigns, matches to the voter files from the firms L2 and TargetSmart, candidates' campaign-finance-based (CF) ideology scores (Bonica, 2014), and candidates'



Figure 2: Partisan composition of city councils and mayoralships over time.

City Housing Permits Data

To study the impact of the partisanship of city leaders on housing policy, we use data on the housing development permits issued by each city in each year from the Census Bureau's Building Permits Survey. The Building Permits Survey is sent to local building permit officials via a mail or online survey and compiled into datasets by the Census Bureau's Manufacturing and Construction Division.¹³ These data contain statistics on new privatelyowned residential construction at the level of individual permit-issuing jurisdictions by year.

We use the city-level annual summary data files that contain total numbers of buildings and units permitted by the type of housing structure — either single-family detached homes or multifamily buildings of two or more housing units — as well as the multifamily proportion of buildings and units that are permitted. We show the over-time variation in these key outcomes variables for all large cities in our target universe in Figure 3.

It is important to note that permitting of multifamily units is not necessarily equivalent to an expansion in the availability of rental units or affordable housing. Multifamily buildings can, of course, be condominiums or market-rate apartments. It is also important to note that permitting of multifamily units is not the only housing policy lever available to a city (e.g., housing subsidy programs, creation of housing trust funds, and creation of inclusionary

partisanship in elections for other offices.

¹²Appendix B describes our sample of elections data in more detail.

¹³These data are inherently limited to new construction for which there was a building permit issued by a government authority. However, the Census Bureau estimates that less than 3% of all privately owned housing units in permit-issuing jurisdictions are built without a permit (U.S. Census Bureau, 2011).



Figure 3: Over-time changes in cities' housing stock, with the total number of buildings permitted and multifamily percent of buildings (top panel), and the total number of units permitted and multifamily percent of units (bottom panel). Both panels use the full dataset of building permits from all cities with a population over 75,000 in 2020.

zoning programs). However, the permitting of multifamily units is an important measure for us for multiple reasons. First, restrictions on building multifamily units have been a key tool of exclusion in housing policy of American cities (Sahn, 2021; Whittemore, 2021). Second, while we cannot observe the share of multifamily units in our data that are marketrate vs. affordable, recent evidence documents that increasing the supply of multifamily units increases a city's housing affordability, even when the newly constructed units are market-rate (Asquith, Mast, and Reed, 2021; Mast, 2021; Li, 2021).

Land Use Regulation

In order to assess the impact of land use regulatory institutions at the city level, we use data on the power of local councils to veto new development. These data are from the Wharton Land Use Regulatory Index (Gyourko, Saiz, and Summers, 2008), a dataset compiled from responses to a survey mailed by researchers in collaboration with the International City/County Management Association's (ICMA). These data were first collected in 2008 and updated in 2018 (Gyourko, Hartley, and Krimmel, 2021), and contain an array of information on the regulatory landscape in cities with regard to land use. Specifically, we use information from this dataset in the most recent year available about whether a city requires the approval of a majority of local councilors, managers, or commissioners in order for any residential land use changes that do not require re-zoning (i.e. "by-right" development). We use this source of data to assess the differential impact of city leaders' partisanship under different regulatory regimes.

Regression Discontinuity Design

We use a regression discontinuity (RD) design to identify the effect of electing city leaders of different parties on fiscal policy.¹⁴ We exploit the fact that a sharp electoral threshold, 50% of the two-party vote share, determines which party wins city council elections.¹⁵ The validity of the RD design depends on the assumption that only the winning candidate — and not the distribution of units' potential outcomes — changes discontinuously at the threshold (Hahn, Todd, and Klaauw, 2001; Lee and Lemieux, 2010). One way to check the validity of this assumption is to examine the density of observations across the threshold with a McCrary test (McCrary, 2008). We conduct this diagnostic test and find a null result for both council and mayoral elections. We also conduct a nonparametric test (Cattaneo, Jansson, and Ma, 2019) and an equivalence test (Hartman, 2021) for the density of observations. All of these tests suggest that the assumption of continuity of potential outcomes is unlikely to be

¹⁴Previous studies in the urban politics literature have similarly used the regression discontinuity design to examine the local incumbency advantage (de Benedictis-Kessner, 2018; Ferreira and Gyourko, 2009; Trounstine, 2011; Warshaw, 2019), and the effects of politicians' characteristics on other policy issues (e.g. Beach and Jones, 2016; Ferreira and Gyourko, 2009; Gerber and Hopkins, 2011; Hopkins and McCabe, 2012; Kirkland, 2021; Kogan, Lavertu, and Peskowitz, 2021; McBrayer and Williams, 2022).

¹⁵In multimember district elections, we compare the bare-winners and bare-losers for the last seat in the race (e.g., in a district with three seats up for election, we compare the votes of the 3rd and 4th placed candidates).

violated, though we cannot rule out some degree of differences in the density of observations across the threshold for mayoral elections.¹⁶ Consistent with the large-scale validation of electoral regression discontinuity (RD) design studies conducted by Eggers et al. (2015), we also observe no significant discontinuities in lagged values of the running variable or other key placebo variables.¹⁷ In order to increase statistical efficiency, we estimate treatment effects on changes in outcomes rather than on levels (Lee and Lemieux, 2010). In order to account for the lag in time between a politician taking office and their ability to influence policy outcomes, our main analyses focus on the difference between housing outcomes in the election year and the average of outcomes measured two and three years after the election.¹⁸

We estimate the effect of electing a Democratic city councilor (or mayor) rather than a Republican councilor (or mayor) based on the "jump" in outcome variables at the threshold. We model the relationship between the assignment and outcome variables with local linear regression, using the default optimal bandwidth options in the **rdrobust** package in R (Calonico, Cattaneo, and Titiunik, 2014*a*).¹⁹ The optimal bandwidth is chosen to minimize mean-square-error (MSE) and confidence intervals are adjusted to account for remaining bias (Calonico, Cattaneo, and Titiunik, 2014*b*; see also Imbens and Kalyanaraman, 2012).²⁰ Since there are often multiple elections in a given year for a particular city's council, we cluster standard errors by city-year in our analyses of councilors, and we cluster by city in our analyses of mayors.²¹

¹⁶The full results from these tests and histograms showing the density of observations across the threshold are presented in Appendix C.

¹⁷These placebo results are shown in Appendix D, which show the effect of electing a Democrat on lagged versions of the running variables and housing outcomes.

¹⁸This strategy enables us to increase statistical power over a strategy using changes in outcomes between the election year and two years after the election by reducing noise in outcomes from individual years (de Benedictis-Kessner and Warshaw, 2016, 2020; Gerber and Hopkins, 2011).

¹⁹In our main analysis, we use the default local linear regression in rdrobust because Calonico, Cattaneo, and Titiunik (2014b) show that local linear regression models perform well in RD designs with optimal bandwidth selection (see also Cattaneo, Idrobo, and Titiunik, 2019, 41-42).

²⁰Our results are robust to this choice of bandwidth, however. We show our effects for other bandwidths than the optimal-MSE one in Appendix G.

²¹We use the 'cluster' option in **rdrobust**. One final complication for our analysis is that while the vast majority of city councils are small (see Figure A2), our dataset is heavily skewed toward the small number of cities with larger legislatures. To address the over-representation of cities with large councils in our dataset, we weight our analyses of the effects of councilor partianship based on the number of councilors

Results

In this section, we examine the effects of partial partial partial point of housing policy using this analytical framework. We assess the effects of political partial partial point the amount of housing permitted and the composition of that housing (single-family vs. multifamily).

Regression Discontinuity Estimates

In order to determine the *causal* relationship between changes in the partisanship of city leaders and city policy, we begin with our estimates using a regression discontinuity design. This strategy restricts our estimand to a *local* average treatment effect (LATE) of partisanship in close elections, rather than the effect of partisanship *overall*. Yet this strategy also enables us to isolate the causal effect of electing a Democratic city councilor or mayor, rather than a Republican, on local housing policy. This design harnesses knowledge from the cities that had a close council or mayoral race between a Democrat and a Republican, which constitutes the vast majority of the cities in our dataset with elections contested by candidates from both parties.²² This includes places in the Sun Belt like Jacksonville, FL, and Tucson, AZ, but also places in the Northeast like Lynn, MA, and in the West like Fresno, CA.

We plot the results analyzing these effects on the number of multifamily housing units permitted in Figure 4, with the Democratic margin in the election plotted along the horizontal axis — showing Democratic victories to the right of zero and Republican victories to the left. Along the vertical axis we plot the change in the natural log of the number of multifamily housing units plus one, with positive values meaning an increase in multifamily housing units over previous years' level of development and negative values meaning a de-

in each city council relative to the average number of councilors, such that elections for larger councils are weighted less heavily than ones for smaller councils. This enables us to interpret our estimates as the effect of elections in the average *city* rather than the average *election*. This approach prevents the handful of cities with very large councils from driving our results. However, the results using unweighted analyses are very similar.

²²See Appendix B.

crease in multifamily units. The trend lines plot local linear regressions weighted using the triangular kernel within the bandwidth selected to minimize mean-squared error (Calonico, Cattaneo, and Titiunik, 2014*b*). The vertical jump between the two lines at the threshold value of zero along the horizontal axis indicates the effect of electing a Democrat rather than a Republican on housing policy.



Figure 4: The effect of partisanship on changes in the logged number of multifamily units permitted in the fiscal years two and three years after an election

The left panel shows that in the average city, electing a Democratic councilor changes the number of multifamily units permitted in the years after an election by about 5% relative to electing a Republican councilor.²³ This effect is statistically insignificant, however. In contrast, electing a Democrat rather than a Republican as mayor leads to an increase in the change in the logged number of multifamily housing units permitted of approximately 0.75 several years after their election, as shown in the right panel of Figure 4. In other words, Democratic mayors increase multifamily housing production by over 70%. This equates to an increase of approximately 100 multifamily housing units per 100,000 capita, as we show in Appendix H. One reason that mayoral partisanship, but not city council partisanship, affects multifamily housing development could be that mayors are typically the ones that

 $^{^{23}}$ These results, as well as the others presented in visual format in this section, are displayed in tabular form in Appendix E.

appoint members of planning commissions and zoning boards of appeal, which usually control variances to existing zoning regulations as well as rezoning.

To further contextualize the magnitude of our estimates, we note that the estimated effect size is small relative to the overall existing housing stock in the typical city. The average city has roughly 15,000 units in multifamily structures per 100,000 in the population.²⁴ But, as captured in the logged outcome estimates, the effect is large relative to the typical numbers of new units permitted in a given year: in our data, cities approve only 163 units per year per 100,000 in the population on average.

How do these effects on multifamily housing production compare to the same effects of city leaders' partisanship on single-family housing production and total housing production? We display the effects of partisanship on these other outcomes in Figure 5. Each point displays the coefficient from the regression discontinuity analysis of partisanship's effect on that outcome, along with its 90% (thick lines) and 95% (thin lines) robust confidence intervals. While the election of Democratic mayors leads to increases in multifamily housing production, it has null or negative effects on the total number and number of single-family buildings and units permitted. The largest effects on multifamily housing production appear for the number of housing units, rather than buildings – suggesting that while Democratic politicians do somewhat increase the *number* of developments in their city, they have larger effects on the *size* of those developments, which leads to increases in the number of units without necessarily increasing the number of buildings. This may reflect the fact that development of new buildings is more dependent on supply-side factors such as the availability of vacant parcels, while the size of those developments is more easily influenced by the political process.

This increase in multifamily housing units alongside a much smaller increase in the total housing units is a compositional change that can also be represented by the proportion of total housing permitted that is multifamily. We plot the effect of electing a Democrat on

²⁴Authors' calculation of 2020 American Community Survey data.



Figure 5: The effect of partianship on changes in housing permitting in the fiscal years two and three years after an election. Thick bars show 90% confidence intervals and thin bars show 95% robust confidence intervals.

the composition of housing permitted – that is, the proportion of total housing permits that are multifamily – in Figure 6. There are large positive effects of mayoral partisanship on the composition of housing units permitted. Electing a Democrat as mayor rather than a Republican leads to an 8 percentage point larger increase in the multifamily proportion of units 2-3 years after their election.²⁵

Given that we identify effects of mayoral partisanship on housing production, we also examined one downstream consequence of increased housing permits: housing prices. This analysis builds on research on housing across the disciplines of economics and urban planning that has consistently identified the effect of building more housing (and more multifamily housing) on the affordability of housing in cities (e.g. Glaeser, Gyourko, and Saks, 2005; Glaeser and Gyourko, 2018). Some recent work finds that a newly constructed residential project, with an average number of units of 165, reduces rent by 6 percent in the surrounding area (Asquith, Mast, and Reed, 2021). This may be driven both by an immediate supply effect and a broader "ripple effect" resulting from the construction of new market-rate units;

²⁵These effects also appear to be enduring for the years 2-4 after the election, which we show in Appendix F, alongside analyses using different time horizon averaging of the outcome variable. We also present robustness checks for these analyses using alternative transformations of the main outcome variables in Appendix H: both a non-logged per capita measure, and the natural log of the outcome plus 0.1 rather than the natural log of the outcome plus 1.



Figure 6: The effect of partial point on changes in type of housing permitted in the fiscal years two and three years after an election. Thick bars show 90% robust confidence intervals and thin bars show 95% robust confidence intervals.

households move into the new units from other – often older – units within the city, expanding the availability of lower-rent units within the city by roughly 45-70 additional units for each 100 newly constructed units, further increasing supply (Mast, 2021). We might therefore expect that our results documenting changes in multi-family housing units would also affect prices.

To assess this, we incorporated data from the Zillow Housing Value Index, a dataset constructed for researchers by Zillow. The index provides a monthly, smoothed, seasonallyadjusted measure of home values (single-family residences and condominiums) at the city level. For our purposes, we collapse the index to the city-by-year level. We then analyzed the effects of city councilors' and mayors' partisanship on the overall housing affordability of cities using these data and the same regression discontinuity design as described earlier. Our analyses in Appendix N provide suggestive evidence that electing a Democrat as mayor leads to a decrease in growth in housing prices, relative to the counterfactual of electing a Republican. However, this analysis is under-powered and should be further examined in future work.

To better understand the conditions under which partial selection in government influences housing policy, we also examined potential institutional moderators of the effects of city councilors' and mayors' partisanship on policy in Appendix K. We assess the different effects of partisanship under strong mayor versus council-manager systems of government, the effects in cities that use district- or ward-based elections for city councilors versus those that use at-large elections, and the effects in cities that use partisan versus nonpartisan ballots in their local elections. We also assess the effects of city councilors' partisanship in cities that have larger or smaller city councils, and in places with closely-divided councils as opposed to councils with larger partisan majorities. In all these cases we observe only small differences in the size of the effects of partisanship.²⁶

Robustness of Main Regression Discontinuity Estimates

Our main result thus far is that the election of a Democratic mayor – but not city council members – leads to an increase in multifamily housing production. In the Appendix, we demonstrate the robustness of this result to different modeling choices. We briefly discuss some of those tests here and summarize them in Table 1.

Appendix	Robustness Check
Appendix F	Different time periods for outcomes
Appendix G	Different bandwidths for RDD
Appendix H	Alternative transformations of non-proportion outcomes
Appendix I	Variety of higher order polynomials for RDD
Appendix J	Randomization inference, 2% bandwidth
Appendix M	Nonparametric Difference-in-Difference Models (PanelMatch)

Table 1: Summary of Robustness Checks in Appendix for Mayoral Results

First, we show that the result is robust to different time horizons (though the effects peak two to three years after the election) and different averaging over subsequent years to construct both the baseline levels and change measures we use as our outcome in Appendix F. Our main results are robust to different bandwidths for the RD model, which we document in Appendix G. As in results reported thus far, we find that Democratic mayors have an impact

²⁶It is important to note that we cannot identify the causal effect of institutions on the size of our main effects due to the few over-time changes in institutional configurations. But the lack of large cross-sectional differences in the main effects implies that institutional variation probably does not affect the impact of mayoral partiasnship on housing policy.

on the logged number of multifamily units and the proportion of multifamily units for a wide array of bandwidths, whereas all outcomes are unresponsive to Democratic councilmembers, regardless of bandwidth. We also present results using alternative transformations of the outcome variables that are not proportions in Appendix H, and find similar results, though we caution that the analyses using non-logged outcomes are subject to over-influence from observations with large outlier values of the outcomes. We also obtain similar results using higher order polynomials for our RD models as well as a simple difference in averages within the optimally-selected bandwidth (i.e., a 0-order polynomial) between cities that elected a Democrat versus those that did not in Appendix I; results are similar, indicating that our main results are not simply an artifact of functional form. In Appendix J, we show that we obtain similar results using local randomization inference in a narrow 2% bandwidth on either side of the discontinuity with 10,000 simulations using the rdlocrand package in R (Cattaneo, Titiunik, and Vazquez-Bare, 2016).

Finally, in Appendix M, we present results using a different research design altogether. Namely, we estimate non-parametric difference-in-differences models using the PanelMatch method (Imai, Kim, and Wang, 2021), which compares units with similar treatment histories (i.e. party control) and similar pre-treatment outcomes (i.e. housing permits) that are "treated" with a Democrat taking control of the mayoral office vs. those that are not treated (i.e. a Republican takes control).²⁷ We prefer the regression discontinuity approach presented in the main text, as it better deals with the endogeneity in the likelihood of electing a Democrat. Despite that, we do ultimately find substantively similar, albeit somewhat smaller, results using the difference-in-differences approach.²⁸

²⁷Specifically, we match using Mahalanobis distance on lagged outcomes in the four years prior to treatment.

²⁸Appendix M also presents results from PanelMatch models assessing the effect of Democratic majority control of city councils.

Regulatory Institutions and Veto Power

We next assess the role of regulatory regimes — that is, the limits on development imposed by allowing city councilors and commissioners to prevent new residential development even when it is allowed under existing zoning laws. Specifically, we use data from the Wharton Land Use Regulatory Index on whether a city requires the approval of a majority of local councilors or commissioners, or of the city manager, in order for any residential projects that do not require re-zoning (i.e. "by-right" development).²⁹ Of the 381 large cities in our elections data, 39% have this rule that requires a majority or supermajority of councilors to approve "by-right" land use changes, essentially limiting new development. This type of rule contributes to the stringent regulatory regimes that increase the price of housing in cities (Gyourko and Krimmel, 2021). We note that this institutional regime is primarily a constraint, and not a positive power: it denotes the ability of city councilors to veto new residential development. We therefore would expect weaker partisan differences on housing outcomes in municipalities with this type of regulation.



Figure 7: The effect of partisanship on changes in the composition of housing permitted in the fiscal years two and three years after an election, divided by the regulatory power afforded to city councils. Thick bars show 90% robust confidence intervals and thin bars show 95% robust confidence intervals.

In Figure 7 we plot the effects of city councilors' and mayors' partisanship on the compo-

²⁹The Wharton Index also reports whether cities require the approval a majority of local councilors or commissioners for dvelopment that *would* deviate from relevant zoning code, but that is true of almost all cities in our sample.

sition of housing permitted under each type of this regulatory regime. These results indicate that when local councilors have veto power over land use changes, this limits the degree to which partisanship influences the housing permits that are issued. Electing a Democrat to the office of city councilor or mayor has no effect on housing in cities where councils have this power. While the partisanship of city councilors still does not cause increases or decreases in the housing permitted when they do not have this veto power, the partisanship of the mayor does have an effect in places without this regulatory institution. Electing a Democrat rather than a Republican as mayor only increases multifamily housing permitted when city councils do not have the ability to veto new development.³⁰ Though the lack of data on over-time variation in this institutional regime prevents us from rigorously assessing whether this moderation is causal (Bansak and Nowacki, 2022), our results suggest that regulatory institutions beyond zoning may influence the growth of housing supply.

Conclusion

Large cities in the United States face a housing affordability crisis. Municipal governments are a crucial thread in the fabric of American democracy and in particularly expensive cities, they may have played some role in the development of this crisis. Furthermore, in cities that are experiencing rapid population growth and are just now seeing the beginnings of the housing crisis, municipal governments may play a role in expanding housing supply and potentially ameliorating increases in housing prices. Investigating the impact of political leaders on housing policy — and how elections influence policy — can help explain contemporary and future housing policy crises.

Assessing the degree to which partian selection functions in city governments is also critical for a broader understanding of democratic functioning in the United States as a whole. While there is a growing body of evidence that the partian composition of local

 $^{^{30}{\}rm We}$ see similar moderation in the effect of partial partial on the number of units permitted, as Figure A25 in Appendix K shows.

governments can affect fiscal policy (de Benedictis-Kessner and Warshaw, 2016, 2020; Gerber and Hopkins, 2011), recent work has suggested limits to the influence of partisanship in local politics (Anzia, 2021, 2022). Some policy arenas may simply not generate disagreement among members of the public or among local politicians (e.g. Thompson, 2020). However, there is no previous study that examines the effects of partisan selection in city councils, and few that assess the role of partisanship in many of the more contentious policy debates that occupy municipal politics. In this paper, we provide a comprehensive assessment of partisanship's effects on the permitting of housing.

We show a variety of evidence that partisanship influences housing policy. Using public opinion data, we show that there are partisan divides in policy views on multifamily housing. Democrats are more supportive of multifamily housing construction. Moreover, we demonstrate that partisan selection in city governments influences housing policy. We find that the partisanship of mayors has large effects on housing policy. When a Democrat is barely elected as mayor rather than a Republican, cities permit more multifamily housing. The partisanship of local politicians can shape subsequent housing outcomes — indicating that partisan divisions among the public in housing policy opinions can translate into differences in local housing policy outcomes due to mayoral elections. We also show preliminary evidence of one important downstream consequence of these increases in multifamily housing: electing a Democrat as mayor may lower housing prices.

We also examine whether the size of these effects varies by the institutional constraints under which local politicians operate. In particular, we find that both the council-manager municipal form of government and nonpartisan ballots, which were instituted as part of the reform movement to insulate city politics from the mores of partisan national politics, do not prevent mayors' partisanship from influencing policy. On the other hand, the city-level zoning rules that give city councilors veto power over land use changes do moderate the effects of mayoral partisanship on housing. When these powers are strong – and allocate more power to city councilors – mayors have less influence over housing policy.

Overall, we demonstrate that city councilors play a nuanced role in the local policy landscape around housing. Despite the fact that housing development is a contentious issue that is often debated in council meetings and over which city councilors are often thought to have control (Einstein, Glick, and Palmer, 2019; Hankinson, 2018; Hankinson and Magazinnik, 2022), the partial point of councilors has no detectable impact on housing policy. City councilors may have few levers by which to influence housing policy. Alternatively, partisanship - while an important feature of politics for other local politicians and on other policy issues - may not structure city councilors' decisions on housing. Instead, city councilors' housing policy decisions may be influenced more by the activity of growth and development interest groups in their cities (Anzia, 2022) rather than their partisanship. Mayors' partisanship, on the other hand, can shift local policy in their ideological direction by a substantively large amount. Yet our analyses show that mayoral partian influence is conditioned by the presence of city councilor veto power. Thus city legislators maintain some ability to affect housing production due to intra-city institutional arrangements. These results suggest that policymaking in cities depends on the distribution of power not just between city and state governments (Palmer et al., 2019), but also *within* city governments.

Our findings extend theories of partian selection and representation to the most common municipal elected office of the city councilor. Moreover, our results extend theories of mayoral partianship to housing policy, arguably the most contentious local policy issue of the current era. Together, these results suggest that theories about the role of partianship and political representation at the local policy should take note of the complex institutional context in each policy area. Housing involves an amalgam of different zoning policies across cities, which combines with different institutional rules surrounding the balance of power among elected and unelected officials in local government. Our findings place the role of these institutions in conversation with theories of representation, and help develop a more holistic understanding of the role of partianship in local policy.

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Supplementary Appendix for "How Partisanship in Cities Influences Housing Policy"

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A Housing Policy Survey

Upon entrance to our survey, potential respondents were given information about the survey and asked for their informed consent. The explanation of the study provided to research participants, with relevant identifying portions of the information redacted, is shown below.

You are invited to participate in a survey about politics and public affairs that is being conducted by [names and universities]. You will be asked to answer a number of questions about national and community affairs. The survey should take approximately 10 minutes to complete. You may be invited to participate in additional follow-up surveys.

Your participation is voluntary. You must be 18 years or older to participate. The only potential risk of this study is a loss of confidentiality, but this is a very small risk. No identifying information other than very general demographic information will be included along with your responses. Taking part in this study will not benefit you directly, but this research may benefit society by improving our understanding of politics and government. If you have read this form and have decided to participate in this study, please understand your participation is voluntary and you have the right to withdraw your consent or discontinue participation at any time without penalty or the loss of benefits to which you are otherwise entitled. The alternative is not to participate. We will not ask for your name or other identifying information. Your individual privacy will be maintained in all published work or public presentations resulting from the study.

If you have any questions, concerns or complaints about this research, its procedures, risks and benefits, please contact [name] at [email]. The [office name] at [university name], at [phone number], can provide further information on your rights as a research participant.

If you consent to take the survey, please click the arrow below to begin.

Participants then were asked a number of policy questions, political questions, and demographic questions, among which was our main question of interest on housing policy, which is shown below.

- 1. Which of the following statements about multifamily housing, such as apartment buildings, comes closer to your view, even if neither is exactly right?
 - Multifamily housing should be allowed everywhere in my city to revitalize neighborhoods and local economies
 - Multifamily housing should be allowed only in dense areas of my city to preserve the character of less populated neighborhoods

B Elections Data Sample

In this section, we provide further details on our elections data. Tables A1 and A2 provide further details on the total elections data gathered as well as those elections used in our descriptive and RDD analyses. The cities in our council and mayoral elections dataset encompass 93% and 99%, respectively, of the population in our target universe of medium and large cities with that type of elections. Moreover, the elections that have a Democratic vote share between 40% and 60%, which roughly approximates the effective sample in many of our RDD analyses, covers 78% (for council elections) and 66% (for mayoral elections) of the population in our target universe overall.

Table A1: Summary of City Council Elections Data Coverage

Subset	N Cities	N Elections	Min Pop.	Max Pop.	Avg. Pop.	Total Pop.	% of Target Uni. Pop.
All cities	19,481		0	8,804,190	10,526	205,058,014	
Medium and large cities (target universe)	476		75,102	8,804,190	224,297	106,765,546	100
Medium and large cities in elections dataset	381	15,520	75,102	8,804,190	259,524	$98,\!878,\!557$	93
Two-party contested elections in dataset	338	3,376	75,604	8,804,190	270,622	91,470,101	86
Two-party close elections in dataset	298	1,503	75,781	8,804,190	281,057	83,755,010	78

Table A2:	Summary	of Mayoral	Elections	Data	Coverage
		•/			0

Subset	N Cities	N Elections	Min Pop.	Max Pop.	Avg. Pop.	Total Pop.	% of Target Uni. Pop.
All cities	19,481		0	8,804,190	10,526	$205,\!058,\!014$	
Medium and large cities	476		75,102	8,804,190	224,297	106,765,546	
Medium and large cities w/ mayoral elections (target universe)	419		75,102	8,804,190	240,204	$100,\!645,\!272$	100
Medium and large cities in elections dataset	396	3,238	75,102	8,804,190	252,594	100,027,292	99
Two-party contested elections in dataset	285	1,045	75,604	8,804,190	282,038	80,380,921	80
Two-party close elections in dataset	218	501	75,644	8,804,190	303,561	$66,\!176,\!369$	66

Figure A1 displays the temporal coverage of these data for both city council (left panel) and mayoral elections (right panel).



Figure A1: Temporal Coverage of Elections Data

Furthermore, Figure A2 shows the size of the legislatures in the cities in our dataset. Typical city councils are fairly small, and the median city in our data has 8 councilors. While councils with fewer than 10 members are typical, some cities — such as New York City — have councils with up to 51 members.



Figure A2: Size of City Legislatures in our Dataset

C McCrary Tests on the Density of Observations

In this appendix we present the results of the McCrary test for the continuity of the density of observations across the 50% vote threshold. These tests replicate the RDD framework but using the density of observations as the outcome. If the density of observations were to have a "jump" in numbers across the threshold, it would suggest a potential violation of the assumption that potential outcomes are continuous at the threshold.

In Table A3 below we present the results of these tests using the number of observations within half-percentage-point bins of voteshare. The coefficient in the second line, indicating the change in the number of observations at the threshold, represents the RDD effect on this outcome. We find a null effect for both city council elections and mayoral elections, suggesting that the continuity assumption is likely to hold in both council and mayoral races.

	Dependent variable:		Dependent variable:
	Number of observations in bin		Number of observations in bin
Voteshare bin	277.922*** (50.605)	Voteshare bin	$79.739^{***} \\ (17.272)$
Voteshare ≥ 0.5	6.042 (4.337)	Voteshare ≥ 0.5	0.514 (1.692)
Voteshare bin \times Voteshare ≥ 0.5	-510.390^{***} (71.567)	Voteshare bin \times Voteshare ≥ 0.5	-143.739^{***} (24.427)
Constant	42.829*** (3.067)	Constant	$ \begin{array}{c} 14.659^{***} \\ (1.196) \end{array} $
$\overline{\text{Observations}} \\ \text{R}^2$	42 0.636	$\begin{array}{c} \text{Observations} \\ \text{R}^2 \end{array}$	48 0.464
Note:	*p<0.1; **p<0.05; ***p<0.01	Note:	*p<0.1; **p<0.05; ***p<0.01

(a) City Council Elections

(b) Mayoral Elections

We also present these results visually in Figure A3, which shows the binned number of

observations both below and above the 50% vote threshold. Visual inspection supports the more formal results shown in Table A3: that there is no discernable effect on the density of observations at the threshold for council elections but a noticeable increase in the density of observations above the threshold for mayoral elections.



Figure A3: Histograms of the number of observations within half percentage-point bins.

However, these tests are subject to a variety of researcher degrees of freedom – in particular, the choice of the size of bin within which to group observations. An alternative check suggested by Cattaneo, Jansson, and Ma (2019) involves conducting a nonparametric test for a discontinuity in the density of the running variable that does not require binning. We present the results from these nonparametric tests, estimated using the R package rddensity, in Table A4 below. Similar to the tests discussed earlier, they indicate no evidence of sorting across the threshold for council elections. However, this nonparametric test does indicate that the difference in the density of observations for mayoral races is statistically distinguishable from zero at the 95% significance threshold, suggesting some evidence of sorting in mayoral elections.

Table A4: Nonparametric Density Tests

(a) City	^v Counci	Council Elections (b)			Iayoral H	Elections
t.statistic	p.value	Effective.N		t.statistic	p.value	Effective.N
0.19	0.85	1543	-	0.35	0.73	568

Finally, others have recently suggested constructing an equivalence test (Hartman and Hidalgo, 2018) based on the density of the forcing variable and calculating inverted p-values based on the null hypothesis of a difference in the density to the left and the right of the cutpoint (Hartman, 2021). We present results using this method in Table A5 below, which show the observed ratio between the density to the left and right of the threshold as well as the equivalence confidence interval and the p-value for the null hypothesis of a jump of greater than 50% in the density across the threshold. This test indicates that the null hypothesis of a substantively important difference in densities can be rejected for council

elections at the 90% confidence level but cannot be rejected for mayoral elections. In both cases, the equivalence confidence interval suggests that the range of differences in density is fairly small in size as well.

Table 4	A5:	Density	Equivalence	Tests
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(a) City Council Elections			()	b) Mayoral Elections	
Observed.Ratio	Equivalence.Confidence.Interval	p.value	Observed.Ratio	Equivalence.Confidence.Interval	p.value
0.97	(0.76, 1.31)	0.01	0.92	(0.62, 1.61)	0.09

D Placebo Tests



Figure A4: Placebo effect of partisanship on lagged democratic voteshare.



Figure A5: Placebo effect of partisanship on pre-treatment # of multifamily units permitted.



Figure A6: Placebo effect of partisanship on pre-treatment ratio of multifamily units.

E Housing Policy RDD Results in Tabular Format

DV	Coef	p-value	BW	Obs
Total buildings, $T+2/3$ Avg	0.04	0.7	11.97	1496
	(-0.18, 0.27)			
Single-family buildings, $T+2/3$ Avg	0.05	0.66	12.63	1554
	(-0.18, 0.28)			
Multi-family buildings, $T+2/3$ Avg	0.03	0.67	12.64	1556
	(-0.25, 0.38)			
Total units, $T+2/3$ Avg	0.03	0.85	10.26	1345
	(-0.26, 0.31)			
Single-family units, $T+2/3$ Avg	0.05	0.66	12.63	1554
	(-0.18, 0.28)			
Multi-family units, $T+2/3$ Avg	0.05	0.76	9.8	1311
	(-0.48, 0.65)			

Table A6: Effect of Councilor Partianship on $\Delta \log(\text{Outcome} + 1)$

Table A7: Effect of Councilor Partisanship on Δ Housing Composition

DV	Coef	p-value	BW	Obs
Multi-family proportion of buildings, $T+2/3$ Avg	-0.01	0.54	14.18	1668
	(-0.04, 0.02)			
Multi-family proportion of units, T+2/3 Avg	-0.03	0.51	10.54	1364
	(-0.11, 0.05)			

DV	Coef	p-value	BW	Obs
Total buildings, $T+2/3$ Avg	-0.03	0.62	9.65	451
	(-0.31, 0.19)			
Single-family buildings, $T+2/3$ Avg	-0.06	0.5	10.45	480
	(-0.38, 0.19)			
Multi-family buildings, $T+2/3$ Avg	0.37	0.04	12.72	531
	(0.02, 0.77)			
Total units, $T+2/3$ Avg	0.11	0.55	8.82	417
	(-0.19, 0.35)			
Single-family units, $T+2/3$ Avg	-0.06	0.5	10.45	480
	(-0.38, 0.19)			
Multi-family units, $T+2/3$ Avg	0.75	0.02	10.18	467
	(0.14, 1.5)			

Table A8: Effect of Mayoral Partisanship on $\Delta \log(\text{Outcome} + 1)$

Table A9: Effect of Mayoral Partisanship on Δ Housing Composition

DV	Coef	p-value	BW	Obs
Multifamily proportion of buildings, T+2/3 Avg	0.03	0.19	12.08	510
	(-0.01, 0.08)			
Multifamily proportion of units, $T+2/3$ Avg	0.08	0.05	12.47	516
	(0, 0.18)			

F Long-run Effects of Partisanship



Figure A7: Long-term effect of partisanship on Δ in logged # of multifamily units permitted.



Figure A8: Long-term effect of partial properties on Δ in multifamily properties.



Figure A9: Effect of partisanship on Δ in logged # of multifamily units permitted averaged over different time horizons.

G Results with Alternative Bandwidths



Figure A10: Effect of partisanship on permitted multi-fam. buildings w/ altern. bandwidths.



Figure A11: Effect of partisanship on multi-fam. units with alternative bandwidths.





H Analyses Using Alternative Transformations of Outcome Variables

A number of our main analyses use the outcome of the natural log of the number of housing units/buildings permitted plus one in order to reduce the influence of large outliers in our outcome variable on our results. The tradeoff of this choice, of course, is that our results using logged outcomes give more influence to data points with smaller values. Though this does not affect our results using proportions (e.g. the multifamily proportion of housing units or buildings), in this section we replicate our main results using two alternative transformations of the outcome variables to test the robustness of our results to the choice of outcome transformation: a non-logged per 100,000 capita measure, and the natural log of the number of housing units/buildings plus 0.1 (rather than +1). The results from these alternative transformations largely corroborate the results presented in the main paper.

In Figure A13 we present the results removing the 5 cities with the largest absolute values of the change in per 100,000 capita multifamily units.³¹ In Figure A14 we also provide the

³¹These cities are Irvine, CA, Nashville, TN, Orlando, FL, Henderson, NV, and Raleigh, NC.

results using the full dataset of non-logged outcome variables. Finally, in Figure A15 we present results of analyses using outcomes measured as the natural log of the outcome + 0.1 rather than the natural log of the outcome + 1, as used in the main manuscript.



Figure A13: The effect of partisanship on changes in type of housing permitted in the fiscal years two and three years after an election.



Figure A14: The effect of partisanship on changes in type of housing permitted in the fiscal years two and three years after an election, including outliers.



Figure A15: The effect of partisanship on changes in type of housing permitted in the fiscal years two and three years after an election.

I Results with Alternative Polynomials

This section shows results using alternative polynomials for the RDD models.



Figure A16: Effect of partisanship on the change in logged multifamily units between the election year and the average of the years two and three years after the election.



Figure A17: Effect of partial properties on the change in the multifamily proportion of buildings between the election year and the average of the years two and three years after the election.



Figure A18: Effect of partial proportion of units between the election year and the average of the years two and three years after the election.

J Results Using Randomization Inference

Table A10: RI Effect of Councilor Partisanship on Δ CBPS Outcomes

DV	Diff. in means	Asymptotic p-value	Obs.	BW
Multi-family units, T+2/3 Avg	-0.06	0.77	337	2
	(-0.45, 0.27)			
Multi-family units, T+2-4 Avg	-0.01	0.97	337	2
	(-0.39, 0.33)			
Multi-family units, Avg. of 2-4 years post-election - 4-yr avg. pre-election	0.17	0.29	337	2
	(-0.09, 0.45)			
Multi-family units, Avg. of 1-4 years post-election - 4-yr avg. pre-election	0.23	0.12	348	2
	(-0.06, 0.48)			

DV	Diff. in means	Asymptotic p-value	Obs.	BW
Multi-family units, T+2/3 Avg	0.77	0	117	2
	(0.26, 1.26)			
Multi-family units, T+2-4 Avg	0.66	0.01	117	2
	(0.15, 1.15)			
Multi-family units, Avg. of 2-4 years post-election - 4-yr avg. pre-election	0.66	0	117	2
	(0.25, 1.05)			
Multi-family units, Avg. of 1-4 years post-election - 4-yr avg. pre-election	0.33	0.1	119	2
	(-0.07, 0.65)			

Table A11: RI Effect of Mayoral Partisanship on Δ CBPS Outcomes

K Moderators

Here, we examine three potential institutional moderators of the effects of city councilors' partial partial partial presence of a strong mayor system (rather than a city manager), the use of at-large versus district elections to elect city councilors, and the use of partial versus nonpartial ballots in local elections. Though we cannot identify the causal effect of any of these institutional configurations on the effects we observe, the cross-sectional differences (and lack of differences that we observe) are still interesting.

In order to assess the impact of different institutions at the city level, we use records of the form of government of cities in our data, their councilor election methods, and whether they use partian ballots for their elections. These data are from the International City/County Management Association's (ICMA) Form of Government surveys.³² We use this source of data to assess the differential impact of electing a Democrat or a Republican on policy in cities with different institutions.

Form of Government

We display the effects of electing a Democrat on types of housing permitted in Figure A19 and on the composition of housing permitted in Figure A20, divided up by form of government.

For both councilors and mayors, the effects of partisanship on housing outcomes appear to be relatively similar in cities with strong mayor and council-manager systems. This suggests that the influence of mayors on housing policy is not confined to cities where they operate as a "strong mayor," and that the lack of influence by city councilors is not only due to their operation in the shadow of "strong" mayors.

³²The ICMA collected these data via a survey sent to city and county government officials every few years from 1974 to 2011 (in 1974, 1981, 1986, 1991, 1996, 2001, 2006, and 2011).



Figure A19: The effect of partial partial partial on changes in housing permitting in the fiscal years two and three years after an election, by form of government.



Figure A20: The effect of partisanship on changes in the composition of housing permitting in the fiscal years two and three years after an election, by form of government.

District and At-Large Elections

We display the effects of electing a Democrat on the type of housing permitted in Figure A21 and on the proportion of housing permitted that is multifamily in Figure A22, divided up by the method by which city councilors in the city are elected. As Hankinson and Magazinnik (2022) argue, district elections (rather than at-large elections) may incentivize city councilors to try and block new housing development in their district, leading to overall decreases in the supply of housing. The effects of partisanship on housing outcomes appear to vary only slightly between cities with different types of councilor elections. City councilors appear to have little influence on housing policy regardless of the method by which the city elects its councilors. Mayors are able to influence the number of multifamily housing units to a larger degree in cities with at-large city council elections, lending suggestive evidence to support the theory that districted council elections may help stymie housing development and the influence of mayors. However, the difference in the sizes of effects is not significant by election method, nor are the results for the multifamily proportion of units in line with these. In addition, very few cities change their method of electing city councilors over the course of our dataset, so we cannot make conclusions about whether this institution has any *causal* effect on the influence of partisanship.



Figure A21: The effect of partial part



Figure A22: The effect of partial part

Partisan Ballots

We display the effects of electing a Democrat on the type of housing permitted in Figure A23 and on the proportion of housing permitted that is multifamily in Figure A24, divided up by the type of ballot used in that city.



Figure A23: The effect of partial part

The effects of partisanship on housing outcomes do not appear to only occur in cities with officially partisan election ballots — and in fact, the effects appear larger in places with nonpartisan ballots. Of course, the majority of cities (63% of elections in our data) hold nonpartisan elections, and very few cities change their ballot form over the course of our dataset, so we cannot make conclusions about whether this institution has any *causal* effect on the influence of partisanship. But these results indicate that officially partisan ballots are not necessary for the partisanship of city leaders to have an effect on policy outcomes.



Figure A24: The effect of partisanship on changes in housing permitting in the fiscal years two and three years after an election, by type of election ballot.

Regulatory Regime

As we show in the main paper, the regulatory regime of cities – the veto powers afforded to city councils over land use development – moderates the effect of partial partial on the multi-family composition of housing permitted. In Figure A25 we show the effects of partial partial permitted by whether or not cities give councils this veto power.



Figure A25: The effect of partisanship on changes in type of housing permitted in the fiscal years two-three years after an election, divided by the regulatory power of city councils.

Partisan Control of Council

In this section, we examine whether the effect of individual legislators is larger when they have the potential to influence partisan control of the city council. Specifically, we analyze whether the effect of partisan selection in city councils on housing is larger in legislatures where, at the time of the election, the partisan majority is small compared to the effect when one party controls a large proportion of the legislature. In closely divided city councils, an additional Democratic legislator could influence majority control of the body, while in more extreme legislatures the partisan majority is unlikely to change. We might therefore expect to observe a larger effect of electing a Democratic legislator rather than a Republican legislator in these more evenly split legislatures.

The top line in Figure A26 indicates the effect of electing a Democratic legislator rather than a Republican legislator in these closely-split legislatures (i.e., where a swing of 1 or 2 seats determines majority control), while the second line indicates this effect in legislatures with three or more seats majority. In the left-panel, we find suggestive evidence that the election of partian legislators may have a larger influence on housing policy when the legislature is closer to evenly split between parties. But none of the results are statistically



Figure A26: The effect of legislative partial on changes in logged per capita expenditures by the partial majority's margin. Thick bars show 90% confidence intervals and thin bars show 95% confidence intervals.

significant, nor are the differences between elections with closely-split and more lopsided councils. Moreover, the right-panel shows that this suggestive effect totally disappears when we focus on the share of housing that is multifamily.

Council Size

Here, we present analyses of the effect of city councilors' partial point on housing by the size of the city council in that city (Figure A27). We find suggestive evidence that electing a Democratic councilor may influence the number of multifamily housing units when those councilors are elected to smaller councils, but these differences are not definitive.



Figure A27: The effect of city councilors' partial partial partial on changes in type of housing permitted in the fiscal years two and three years after an election, by the size of the city council.

L Effects of Partisanship on Alternative Housing Outcomes

In this section, we assess the effects of politicians' partial policy on several alternative housing policy-related outcomes: the eviction rate, spending by local governments of HUD funds (used for affordable housing), and the number of new low-income housing tax credit (LIHTC)-subsidized housing units developed. Our results indicate no consistent effects of city councilor or mayoral partial partial policy on these outcomes, though they show suggestive effects of mayoral partial partial policy.



(a) City Council Elections

(b) Mayoral Elections

Figure A28: Effects of partisanship on the eviction rate.



(a) City Council Elections

(b) Mayoral Elections

Figure A29: Effects of partisanship on total HUD spending.



(a) City Council Elections

(b) Mayoral Elections

Figure A30: Effects of partisanship on number of units of LIHTC-subsidized housing units built.

M Difference-in-Differences Effect of Partisanship on Housing Permits

In the main manuscript, we present RDD results for the effect of city councilor and mayoral partisanship on housing permits. Below, we present results from the PanelMatch method (Imai, Kim, and Wang, 2021), which compares units with similar treatment histories (i.e. party control) that are "treated" with a Democrat taking control of the mayoral office vs. those that are not treated (i.e. a Republican takes control). Figure A31 shows the effect of Democratic mayoral control on several of our primary housing outcome measures. Consistent with our main results, these analyses suggest that the election of a Democratic mayor increases the logged number of multifamily housing units permitted, as well as the multifamily proportion of total units permitted in the years after they take power.

In addition, we assess the effect of majority partian control of city councils on housing using a similar DID framework. These results are displayed below in Figure A32, and indicate that there is little effect of Democratic majority control of city councils on housing outcomes.



Figure A31: Effects of mayoral partial partial on housing outcomes using PanelMatch. Thick bars show 90% confidence intervals and thin bars show 95% confidence intervals.



Figure A32: Effects of city council Democratic control on housing outcomes using PanelMatch. Thick bars show 90% confidence intervals and thin bars show 95% confidence intervals.

N Consequences for Affordability

Research on housing across the disciplines of economics and urban planning has consistently identified the effect of building more housing (and more multifamily housing) on the affordability of housing in cities (e.g. Glaeser, Gyourko, and Saks, 2005; Glaeser and Gyourko, 2018). Given that we identify effects of mayoral partisanship on housing production, we also examined this downstream consequence of increased housing permits. To do so, we incorporated data from the Zillow Housing Value Index, a dataset constructed for researchers by Zillow. The index provides a monthly, smoothed, seasonally-adjusted measure of home values (single-family residences and condominiums) at the city level. For our purposes, we collapse the index to the city-by-year level. We then analyzed the effects of city councilors' and mayors' partisanship on the overall housing affordability of cities using these data and the same regression discontinuity design as described earlier. Our analyses provide suggestive evidence that electing a Democrat rather than a Republican as mayor appears to lead to approximately 7% lower housing values in the 2-3 years following their election. It is important to note, however, that these analyses are underpowered.³³



Figure A33: Effects of partisanship on the value of residential housing, as measured by the Zillow Housing Value Index (ZHVI). The ZHVI represents the average value of homes sold in each city in each year. Bars show 90% (thick lines) and 95% (thin lines) robust confidence intervals.

Furthermore, we analyze the robustness of these results on housing prices by varying both the bandwidth and the order of the polynomial used in the RDD. The results of these robustness checks are presented in Figure A34 and Figure A35.

³³In Appendix L we examine several potential causes of decreases in housing prices. Our results in that section suggest that electing a Democrat as mayor may lead to small increases in HUD spending and small increases in new LIHTC-subsidized units developed, which combined with the effects of market-rate multifamily development may explain the decreases in housing prices. In addition, we also examine another possible outcome of increased development: evictions. Our results indicate little influence of mayoral partisanship on evictions.



Figure A34: Effect of partisanship on logged ZHVI using alternative bandwidths. Bars show 95% robust confidence intervals, which are colored blue if they overlap with zero and red if they do not.



Figure A35: Effect of partianship on the change in ZHVI between the election year and the average of the years two and three years after the election using alternative polynomials. Bars show 95% robust confidence intervals.