

**John F. Kennedy School of Government  
Harvard University  
Faculty Research Working Papers Series**

**Is Trade Good or Bad for the Environment?  
Sorting Out the Causality**

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September 2003**

**RWP03-038**

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# Is Trade Good or Bad for the Environment? Sorting Out the Causality \*

Draft: September 15, 2003, for KSG WP

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## Abstract

We seek to contribute to the debate over globalization and the environment by asking: What is the effect of trade on a country's environment, for a given level of GDP? We take specific account of the endogeneity of trade, using exogenous geographic determinants of trade as instrumental variables. We find that trade tends to reduce three measures of air pollution. Statistical significance is high for concentrations of SO<sub>2</sub>, moderate for NO<sub>2</sub>, and lacking for particulate matter. While results for other environmental measures are not as encouraging, there is little evidence that trade has a detrimental effect on the environment.

Keywords: Openness, Growth, pollution, Kuznets, instrumental, variable, causality, simultaneity.

JEL Classification: F18

\* The authors thank for useful comments: Bill Clark, Judy Dean, Dan Esty, Don Fullerton, Arik Levinson, Edward Parson, Rob Stavins, M. Scott Taylor, Jeffrey Williamson, and participants at the NBER Environmental Economics program meeting, Harvard Environmental Economics seminar, Harvard International Economics seminar, and KSG Faculty Lunch. Frankel acknowledges support from the Savitz Family Fund for Environmental and Natural Resource Policy, and research assistance from Anne Lebrun. A current version of the paper and the data set are available at Rose's website. This is a condensed version of a working paper that has the same title, more results, and more references, which is available as NBER working paper #9201 and at both authors' websites. It includes appendix tables for two pollution haven tests: interaction between openness and income or capital/labor ratios. (These results are to be omitted from the published version.) The NBER working paper used a spline for income, whereas this revision uses the quadratic form.

# Is Trade Good or Bad for the Environment? Sorting Out the Causality

## 1: Introduction

Opponents of globalization often fear the adverse effects of trade on environmental quality. Should they? In this short empirical paper, we use cross-country data to show that there is little evidence that openness increases air pollution, holding other things (such as income) constant. Our modest and incremental contribution to the literature is to take special care to account for the fact that income, trade, and environmental quality are determined simultaneously.<sup>1</sup>

The simultaneity issue is potentially important. As Figure 1 shows, a rough inverse correlation between SO<sub>2</sub> concentrations and trade is visible.<sup>2</sup> Eiras and Schaeffer (2001, p. 4), for example, find: “In countries with an open economy, the average environmental sustainability score is more than 30 percent higher than the scores of countries with moderately open economies, and almost twice as high as those of countries with closed economies.” Does this mean that trade is good for the environment? Not necessarily. Causality could run in other directions. The observed correlation might be a result of the Porter hypothesis -- which claims that environmental regulation stimulates productivity -- together with the positive effect of income on trade. Or it might be because democracy leads to higher levels of environmental regulation, and democracy is causally intertwined with income and trade.

The central focus of the paper is to estimate the effect of trade on the environment *for a given level of income per capita*. This is an interesting question for two reasons. First, it is the most fundamental question for policy. If it is established that trade has an adverse effect on the environment solely because openness raise countries' incomes, which in turn damages the

environment, few would conclude that we should try to turn back the clock on globalization. Not many would choose deliberate self-impoverishment as a means to a clean environment.

The question is also interesting because, although the topic is the subject of a rapidly growing area of research, the answer is not settled. Indeed, the effect of trade on the environment is theoretically ambiguous.<sup>3</sup>

Many believe that openness harms the environment. Most widely discussed is the *race to the bottom hypothesis*, which says that open countries in general adopt looser standards of environmental regulation, out of fear of a loss in international competitiveness. Alternatively, poor open countries may act as pollution havens, adopting lax environmental standards to attract multinational corporations and export pollution-intensive goods.<sup>4</sup>

But less widely recognized is the possibility of an effect in the opposite direction, which we call the *gains from trade hypothesis*. If trade raises income, it allows countries to attain more of what they want, which includes environmental goods as well as more conventional output. Openness could have a positive effect on environmental quality (even for a given level of GDP per capita) for a number of reasons. First, trade can spur managerial and technological innovation, which can have positive effects on both the economy and the environment. Second, multinational corporations tend to bring clean state-of-the-art production techniques from high-standard source countries of origin to host countries. Third is the international ratcheting up of environmental standards through heightened public awareness.<sup>5</sup> While some environmental gains might tend to occur with any increase in income, whether taking place in an open economy or not, others may be more likely when associated with international trade and investment.<sup>6</sup> Whether the race to the bottom effect in practice dominates the gains from trade effect is an empirical question.<sup>7</sup>

Our paper is part of a larger literature; a number of studies have sought to isolate the independent effect of openness on the environment. In addition to those already mentioned, Lucas, et al. (1992), study the toxic intensity implied by the composition of manufacturing output in a sample of 80 countries, and find that a high degree of trade distorting policies increased pollution in rapidly growing countries.<sup>8</sup> Harbaugh, Levinson, and Wilson (2000) report in passing a beneficial effect of trade on the environment, after controlling for income. Dean (2002) finds a detrimental direct of liberalization for a given level of income, via the terms of trade, though this is outweighed by a beneficial indirect effect via income.<sup>9</sup> Antweiler, Copeland and Taylor (2001) and Copeland and Taylor (2003) represent an extensive body of theory and empirical research explicitly focused on the effects of trade on the environment. They conclude that trade liberalization that raises the scale of economic activity by 1 per cent works to raise SO<sub>2</sub> concentrations by ¼ to ½ % via the scale channel, but that the accompanying technique channel reduces concentrations by 1 ¼ to 1 ½%, so that the overall effect is beneficial.

Antweiler, et al, point out that endogeneity could be a potential weakness of their work; and a number of authors have sought to address some aspects of endogeneity.<sup>10</sup> But the existing research does not directly address the problem that trade may be simultaneously determined with income and environmental outcomes. Allowing for the endogeneity of trade and income is the main new contribution of this paper.

## **2: Methodology**

We turn directly to the empirics.

## Equation to be estimated

We estimate the following cross-country equation:

$$\begin{aligned} EnvDam_i = & \varphi_0 + \varphi_1 \ln(y / pop_1)_{90,i} + \varphi_2 \ln(y / pop_2)_{90,i}^2 \\ & + \beta([X + M] / Y)_{90,i} + \varphi_3 (Polity)_{90,i} + \varphi_4 \ln(LandArea / pop)_{90,i} + e_i \end{aligned}$$

where:

- $EnvDam_i$  is one of three measures of environmental damage for country  $i$ ,
- $\{\varphi_i\}$  is a set of control coefficients,
- $\ln(y/pop)_{90,i}$  represents the natural logarithm of 1990 real GDP per capita for country  $i$ ,
- $(X + M)/Y$  represents the ratio of nominal exports and imports to GDP (“openness”),
- Polity is a measure of how democratic (vs. autocratic) is the structure of the government,
- $LandArea/pop$  is a measure of per capita land area, and
- $e$  is a residual representing other causes of environmental damage.

The coefficient of interest to us is  $\beta$ , the partial effect of openness on environmental degradation.

Income plays a strong role in determining environmental outcomes. We incorporate into our analysis -- without relying on -- the “environmental Kuznets curve” (EKC). This is a U-shaped relationship between income per capita and certain types of pollution, brought to attention by the World Bank (1992) and Grossman and Krueger (1993, 1995). Growth increases air and water pollution at the initial stages of industrialization, but later on can reduce pollution given the right institutions, as countries become rich enough to pay to clean up their environments. The EKC hypothesis predicts that the coefficient on the squared income term is negative, so that the pollution curve eventually turns down.<sup>11</sup>

The market does not address externalities left to itself. Higher income is unlikely to result in an improved environmental regulation absent appropriate political institutions. Thus it is important to control also for the latter, which we do by including *polity* in our equation.

### **Addressing endogeneity**

The endogeneity of trade is a familiar problem from the empirical literature on openness and growth.<sup>12</sup> What is needed is a good instrumental variable, which is exogenous yet highly correlated with trade. The gravity model of bilateral trade offers a solution. It states that trade between a pair of countries is determined, positively, by country size (GDP, population, and land area) and, negatively, by distance between the countries in question (physical distance as well as cultural distance in the form of, e.g., difference languages). Geographical variables are plausibly exogenous. Yet when aggregated across all bilateral trading partners, these variables are highly correlated with a country's overall trade, and thus make good instrumental variables, as first noted by Frankel and Romer (1999). Thus we construct an instrumental variable for openness by aggregating up across a country's partners the prediction of a gravity equation that explains trade with distance, population, language, land border, land area, and landlocked status.

We use a cross-country approach, thus choosing not to follow Grossman and Krueger (1993) and Antweiler et al (2001) in using panel data. We realize that a pure cross-section approach means that we cannot control for unobservable heterogeneity. But our key instrument is driven by cross-country geographical variation, which does not change over time, so there seems little advantage for us in a panel study.

Income per capita too is endogenous. Both trade and environmental regulation may affect income.<sup>14</sup> We thus use a second set of instrumental variables for income, taken from the

growth literature. These include: lagged income (thus we incorporate the conditional convergence hypothesis), population size, and rates of investment and human capital formation (the factor accumulation variables familiar from neoclassical growth equations).

## Data

We focus on results for three 1990 measures of air pollution, all measured as concentrations in micrograms per cubic meter (simply averaged across a country's measuring stations and cities, in cases where more than one observation was available):

- SO<sub>2</sub>: mean sulphur dioxide,
- NO<sub>2</sub>: mean nitrogen dioxide, and
- PM: mean total suspended particulate matter.

We have also looked at four other measures of environmental quality:

- CO<sub>2</sub>: Industrial carbon dioxide emissions per capita, in metric tons
- Deforestation: average annual percentage change, 1990-95
- Energy depletion: "genuine savings" as a percentage of GDP<sup>15</sup>, and
- Rural Clean Water Access: as percentage of rural population, 1990-1996.

Of these seven, the three measures of local air pollution – SO<sub>2</sub>, NO<sub>2</sub>, and particulates – are the most relevant. CO<sub>2</sub> is a purely global externality, and unlikely to be addressed by regulation at the national level. Deforestation and energy depletion are not measures of pollution, and measuring them involves some serious problems of composition and data reliability, as does water access. Still, it seems worthwhile to look as well at these broader measures of environmental quality.

Per capita income is defined as 1990 GDP per capita (measured in real PPP-adjusted dollars), taken from the Penn World Table 5.6. The Penn World Table also supplies our measure



of openness. Polity ranges from -10 (“strongly autocratic”) to +10 (strongly democratic), and is taken from the Polity IV project. Land area is taken from the CIA’s website and is intended to allow for the likelihood that population density leads to environmental degradation (for a given level of per capita income). Descriptive statistics are included in an appendix table, and simple scatterplots are portrayed in the appendix figure.

### **3: Results**

Table 1 reports our key estimation results, where the dependent variable is represented in turn by the three measures of air pollution. The three columns at the left of the table are the OLS estimates, while the IV estimates are on the right.

The estimated effect of the polity variable on pollution is always negative, suggesting that improved governance has a beneficial effect. It is generally significant statistically. The same is true of land area per capita, evidence that population density has an adverse effect on pollution.

Of greater interest is the relationship with per capita income. The estimated coefficient on the quadratic term is negative for all three measures of air pollution, confirming the EKC hypothesis: after a certain point (recorded at the bottom of the table as “income peak”), growth reduces these environmental indicators. Statistically, it is highly significant in the case of SO<sub>2</sub> and NO<sub>2</sub>, and moderately so in the case of PM.

Our central interest is  $\beta$ , the coefficient on openness. The OLS estimate is negative for all three kinds of air pollution – insignificantly so for PM, moderately significant for NO<sub>2</sub>, and highly significant for SO<sub>2</sub>. Apparently any adverse “race to the bottom” effect on air pollution is outweighed by a positive “gains from trade” effect.

The main contribution of this paper is to address the possibility that these apparent effects may be the spurious results of simultaneity. The right part of the table reports instrumental

variables estimates, where the gravity-derived prediction of openness is the instrument for trade and the factor accumulation variables are the instruments for income. The IV results are generally similar to the OLS results, though with somewhat diminished significance levels in some cases. The EKC is still there for all three pollutants, and the coefficient on openness is negative for all three pollution measures. As in the OLS results, statistical significance is high for SO<sub>2</sub>, moderate for NO<sub>2</sub>, and lacking for particulates.

As an alternative to our quadratic functional form for the EKC, we have also tried a three-segment spline, with “knots” at the .33 and .66 percentiles of the logarithm of per capita income. Results are comparable, and are reported in the working paper.<sup>16</sup>

### **Results for Other Environmental Measures**

Air pollution is only one kind of measure of environmental quality. We also produced analogous estimates for other measures of environmental degradation.

Table 2 reports our OLS and IV estimates of  $\beta$  for carbon dioxide, deforestation, energy depletion, and access to clean water.<sup>17</sup>

Beneficial OLS effects of openness show up only for energy depletion and clean water access (an increase in clean water access indicates a beneficial environmental effect, the reverse of the other six indicators), and are of borderline statistical significance. The case that would give an environmentalist the greatest concern is CO<sub>2</sub>. The coefficient on openness is positive and of moderate significance. This result could be viewed as one piece of evidence supporting the idea that the free-rider problem inhibits individual countries from curbing emissions of greenhouse gases on their own.<sup>18</sup> CO<sub>2</sub> is a purely global externality, so there is no reason to expect individual countries to address it without some mechanism of international cooperation.

When instrumental variables are used, the detrimental effect of openness on carbon dioxide emissions loses all significance, while the apparently beneficial effect on energy depletion becomes significant at the 10% level. On the other hand, the beneficial OLS effect on water access disappears. Evidently the use of instrumental variables to correct for simultaneity can make an important difference to some results.

To summarize: the results are generally supportive of the environmental Kuznets curve, and of the positive effect of democracy on environmental quality. More importantly, there is some evidence that openness reduces air pollution and little evidence that openness causes significant environmental degradation, other things equal. The most important exception is carbon dioxide.

### **Do Some Countries Have a “Comparative Advantage” in Pollution?**

We also test the “pollution haven” hypothesis according to which economic integration results in some open countries exporting pollution to others, even if there is no systematic effect on the world environment in the aggregate.

One version of the hypothesis is that open countries that have a particularly high demand for environmental quality – rich countries – specialize in products that can be produced cleanly, letting poor open countries produce and sell the products that require pollution. This hypothesis can be readily tested by adding the interaction of openness and income per capita to our equation. If rich countries take advantage of trade by transferring the location of pollution-creating activities to poor countries, the interaction between openness and income should have a negative effect on pollution. When we tried this, the coefficient on the interactive term was insignificant for most of the seven environmental measures. The exceptions are particulates and

SO<sub>2</sub>. With either OLS or IV estimation, openness interacted with income has a positive effect on these two types of pollution, opposite of that predicted by the standard pollution haven hypothesis.<sup>19</sup>

A second version of the pollution haven hypothesis is that countries endowed with a high supply of environmental quality – e.g., those with high land area per capita – become pollution havens, exporting dirty goods to more densely populated countries. We tested this by adding the product of openness and land area per capita. Again, signs were divided between negative and positive, and the coefficients were usually insignificant. The only two cases with significant interaction coefficients (IV for particulates, and OLS for CO<sub>2</sub>) have the “wrong” sign, suggesting that more sparsely populated countries have lower emissions than they otherwise would, not higher. Again, there is no evidence supporting the pollution haven hypothesis.

A third possible source of “comparative advantage” derives from traditional trade theory. If some countries have a comparative advantage in capital-intensive sectors such as mining or heavy manufacturing, and these sectors produce comparatively more pollution, then trade may lead to an increase in pollution among the capital-endowed countries and a decrease among the labor-endowed countries.<sup>20</sup> We tested this version by including interactive terms defined as openness times the country’s capital/labor ratio. The signs are mixed, and standard errors large; the interactive term is not statistically significant.

To summarize: there is no evidence that poor, land-abundant, or capital-abundant countries use trade to exploit a comparative advantage in pollution. The only cases where the coefficient on the interactive term appears significant are of the wrong sign. The details of estimates for all three versions of the pollution haven hypothesis are available in the working paper version of this paper.

#### **4: Conclusions**

Trade can have several sorts of effects on the environment. In this short paper, we have modeled the effect of trade on the environment, controlling for income and other relevant factors. The primary contribution of the paper is to address the endogeneity of income and especially trade, the latter by means of instrumental variables drawn from the gravity model. While the use of instrumental variables did not radically reverse the results of earlier OLS studies, it could have; and it did make a substantive difference to the estimates in some cases.

We have found that trade appears to have a beneficial effect on some measures of environmental quality, though not all, *ceteris paribus*. The effect is particularly beneficial for some measures of air pollution, such as SO<sub>2</sub>. Our examination of seven different measures of environmental quality provides little evidence that trade has a detrimental effect overall. We reject the hypothesis of an international race to the bottom driven by trade. There is also no evidence for the pollution haven hypothesis, which claims that trade encourages some countries to specialize in dirtier environments.

Other evidence shows that trade promotes economic growth.<sup>21</sup> Thus trade also has an indirect effect on the environment. Given the environmental Kuznets curve, at low levels of income this effect increases pollution, but at high levels reduces it.

The major example where trade and growth might have the detrimental effects feared by environmentalists is carbon dioxide. Greenhouse gases are global externalities, and there is no reason to expect individual countries to be able to address them in the absence of an international agreement.

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**Table 1: Determinants of Air Pollution Concentrations**

	OLS	OLS	OLS	IV	IV	IV
	NO <sub>2</sub>	SO <sub>2</sub>	PM	NO <sub>2</sub>	SO <sub>2</sub>	PM
<b>Trade / GDP</b>	-.29 (.17)	-.31 (.08)	-.37 (.34)	-.33 (.19)	-.23 (.10)	-.31 (.41)
<b>Log real GDP per capita</b>	409 (122)	287 (119)	567 (336)	461 (199)	296 (140)	681 (412)
<b>Log real GDP p/c squared</b>	-22.8 (6.9)	-16.6 (6.8)	-35.6 (19.1)	-25.6 (10.9)	-17.1 (7.7)	-42.0 (23.2)
<b>Polity</b>	-3.20 (1.47)	-6.58 (2.05)	-6.70 (3.42)	-3.77 (1.37)	-6.41 (2.27)	-7.78 (4.07)
<b>Log of Area per capita</b>	-5.94 (5.93)	-2.92 (1.39)	-13.0 (6.29)	-6.14 (6.43)	-1.54 (1.96)	-12.6 (6.84)
<b>Observations</b>	36	41	38	35	40	37
<b>R<sup>2</sup></b>	.16	.68	.62	.18	.67	.63
<b>Income Peak</b>	\$7665	\$5770	\$2882	\$8015	\$5637	\$3353

Cross-country estimation across countries in 1990.

(Robust standard errors in parentheses.)

Regressands are averages per cubic meter. Intercept included but not reported.

Instrument for trade constructed by aggregating predicted bilateral gravity equation of trade on: distance, population, area, and dummies for language, land border, and landlocked status. Instruments for income (and square) constructed from regression of income on: lagged income, population, openness, investment, population growth, and primary and second school enrollments.



**Table 2: Effect of Openness on her Types of Environmental Degradation**

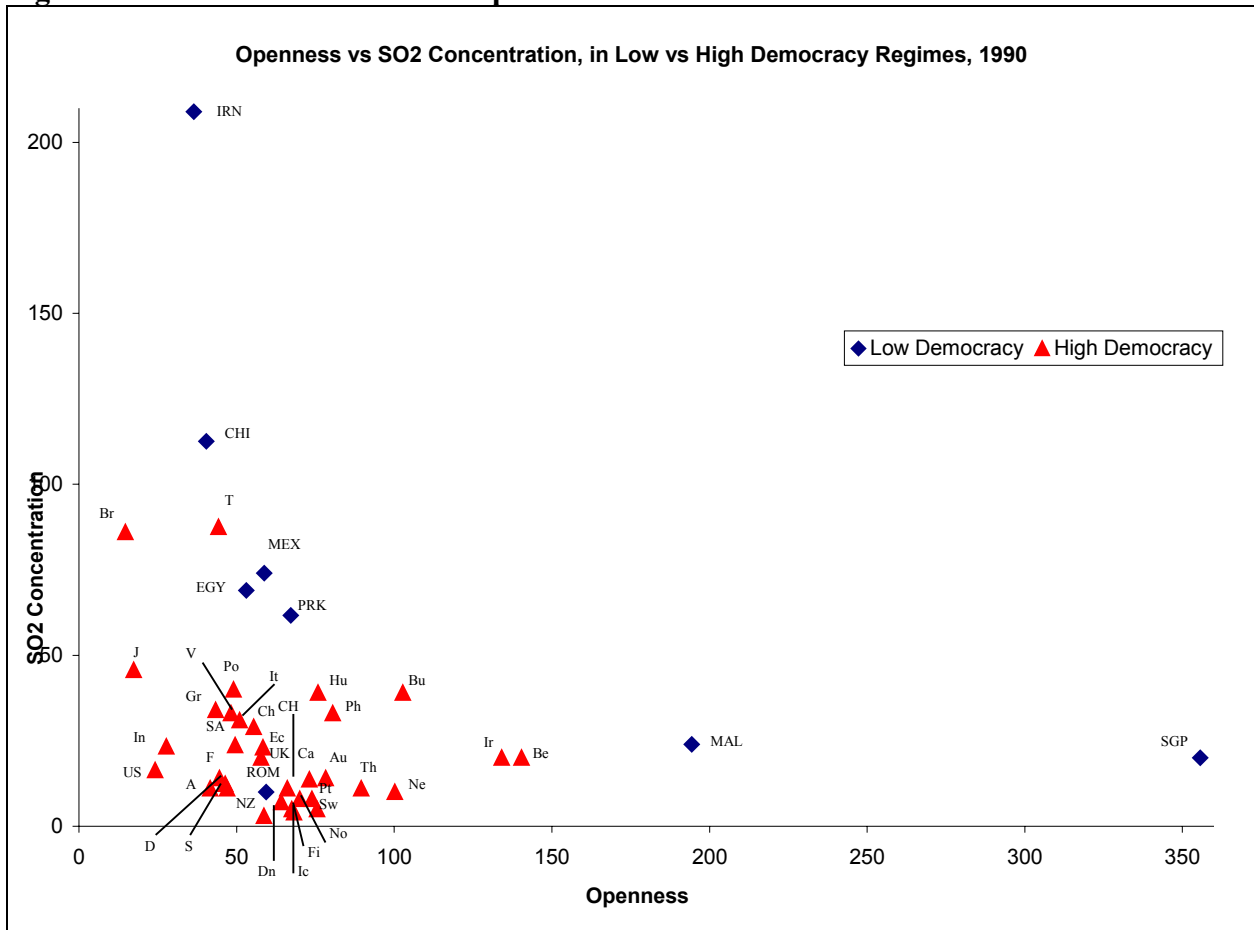
	<b>OLS</b>	<b>IV</b>
<b>CO<sub>2</sub></b>	.016 (.008)	.000 (.010)
<b>Deforestation</b>	.002 (.003)	.001 (.004)
<b>Energy Depletion</b>	-.014 (.009)	-.034 (.020)
<b>Rural Clean Water Access</b>	.111 (.078)	-.067 (.266)

Estimation across countries in 1990.

(Robust standard errors in parentheses.)

Income, income squared, polity score, log area per capita, and intercept were included in the regression, but are not reported here, to save space.

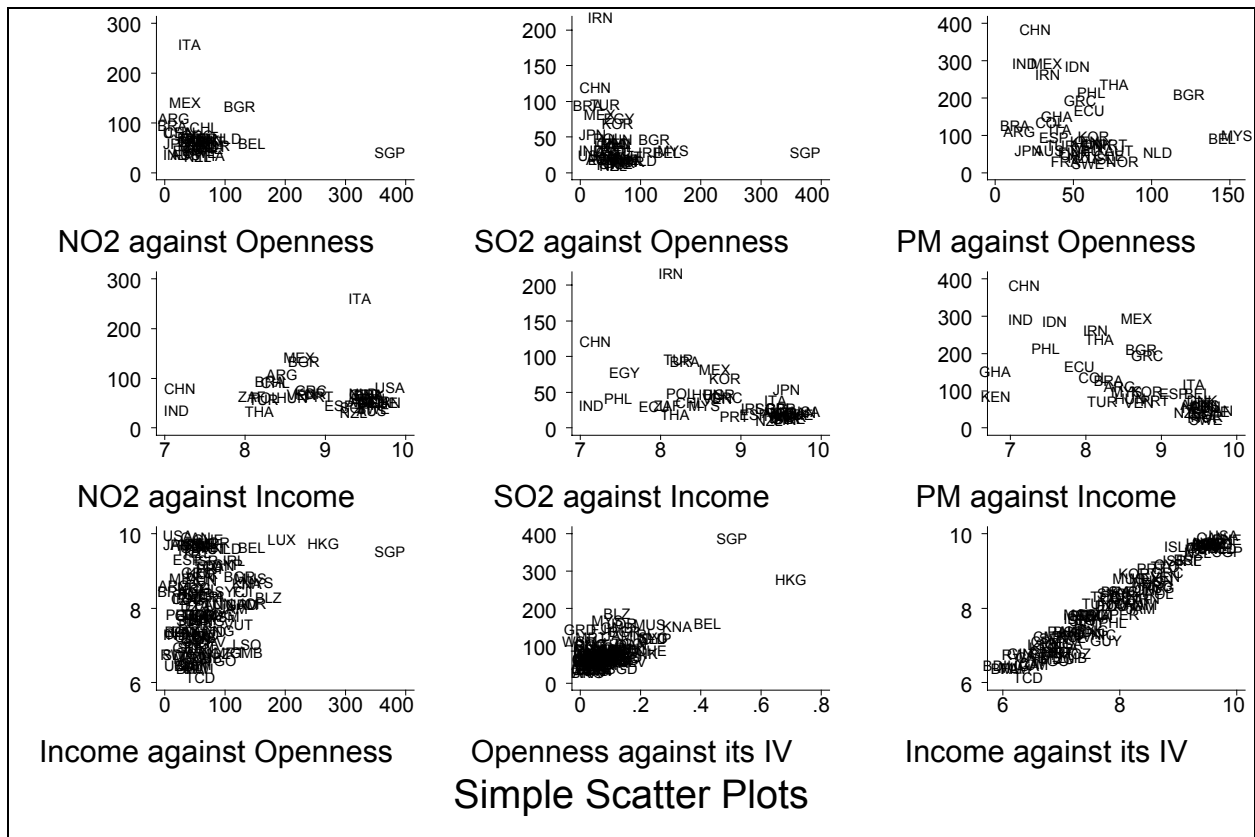
Figure 1: SO2 versus measure of openness to trade



 **Table A1: Descriptive Statistics**

	<b>Obs.</b>	<b>Mean</b>	<b>Standard Deviation</b>	<b>Min</b>	<b>Max</b>
<b>NO<sub>2</sub></b>	42	55.4	39.5	5	248
<b>SO<sub>2</sub></b>	48	31.9	35.9	1	209
<b>Particulate Matter</b>	44	103.1	84.8	9	368.9
<b>CO<sub>2</sub></b>	147	4.1	6.0	.0	31.3
<b>Deforestation</b>	137	.5	1.2	-2.7	7.2
<b>Energy Depletion</b>	144	4.7	12.7	0	104.3
<b>Rural Clean Water Access</b>	70	7.5	5.5	.2	26.9
<b>Trade / GDP</b>	113	73	49	13	373
<b>Log real GDP per capita</b>	113	8.0	1.1	6.0	9.8
<b>Polity</b>	133	.8	7.7	-10	10
<b>Log of Area per capita</b>	112	3.0	1.5	-1.8	6.6

**Figure**  **Simple Scatterplots of data on pollution and trade**



**Appendix tables**

**Results for interaction of income and trade (with income in quadratic form).**

**cited in published paper, but not reported there**

```

. *
. * This differs from pa4c in that interactions of openness with income p/c
. * and land area are added (the interaction with K/L is in pa4ac)
. *
. * This STATA program estimates a set of equations for:
. * a) income;
. * b) openness; and
. * c) pollution.
. *
. * This program uses IV and OLS methods.
. *
. * The new baseline uses quadratics, not a spline for income.
. *

. *
. * First interaction: with GDP p/c
. *
. * Add the interactions to the benchmark equation . OLS then IV.
. *
. * CO2
. *
. reg co2perc inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors                Number of obs =    100
                                                       F( 6, 93) = 40.99
                                                       Prob > F      = 0.0000
                                                       R-squared    = 0.7659
                                                       Root MSE    = 2.3632

-----+-----
      co2perc |           Coef.   Robust
              |           Std. Err.   t    P>|t|
-----+-----
      inc | -17.54647   4.202776   -4.17  0.000
      incsq |  1.247794   .2689828    4.64  0.000
      pwtopen | -.1171033   .0728187   -1.61  0.111
      intery |  .0152322   .0089684    1.70  0.093
      polity | -.0021871   .0295169   -0.07  0.941
      lareapc |  .2897612   .1886449    1.54  0.128
      _cons |  61.60089   16.40485    3.76  0.000

-----+-----
. test inc incsq      ( 1)  inc = 0      ( 2)  incsq = 0
      F( 2, 93) = 14.32                Prob > F = 0.0000

. * Income Peak      1131.1588

```

```

ivreg co2perc (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

```

```

IV (2SLS) regression with robust standard errors                Number of obs =    96
                                                       F( 6, 89) = 24.50
                                                       Prob > F      = 0.0000
                                                       R-squared    = 0.5760
                                                       Root MSE    = 2.8579

```

co2perc	Coef.	Robust Std. Err.	t	P> t
inc	-20.42214	8.420988	-2.43	0.017

```

      incsq | 1.688672   .7398712    2.28   0.025
    pwtopen | .4165342   .5465228    0.76   0.448
      intery | -.0468285   .0616196   -0.76   0.449
      polity | -.136495    .1089114   -1.25   0.213
    lareapc | -.4210369   .7030166   -0.60   0.551
      _cons | 56.35638   21.74199    2.59   0.011
-----+-----
Instrumented:  inc incsq pwtopen
Instruments:  intery polity lareapc elhsfs incf incfsq
-----+-----
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
      F( 2,      89) =      3.06   Prob > F =      0.0518

. * Income Peak      422.7599

. *
. * Deforestation
. *
. reg defp inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors                                Number of obs =      96
                                                                    F( 6,      89) =      7.65
                                                                    Prob > F      =      0.0000
                                                                    R-squared     =      0.2538
                                                                    Root MSE     =      1.1267
-----+-----
      |           |           |           |           |           |
      defp |           |           |           |           |           |
-----+-----
      inc | 4.379583   1.332433    3.29   0.001
      incsq | -.3014252   .0816164   -3.69   0.000
    pwtopen | .0268688   .0288845    0.93   0.355
      intery | -.0028982   .0031489   -0.92   0.360
      polity | .0330982   .0294806    1.12   0.265
    lareapc | -.1405813   .0881114   -1.60   0.114
      _cons | -14.57958   5.434964   -2.68   0.009
-----+-----
. test inc incsq      ( 1) inc = 0      ( 2) incsq = 0
      F( 2,      89) =      9.1   Prob > F =      0.0002

. * Income Peak      1429.0891

ivreg defp (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

IV (2SLS) regression with robust standard errors                    Number of obs =      92
                                                                    F( 6,      85) =      2.98
                                                                    Prob > F      =      0.0108
                                                                    R-squared     =      .
                                                                    Root MSE     =      1.958
-----+-----
      |           |           |           |           |           |
      defp |           |           |           |           |           |
-----+-----
      inc | .9473862   11.67173    0.08   0.935
      incsq | .0808805   1.134537    0.07   0.943
    pwtopen | .4140609   .9814517    0.42   0.674
      intery | -.0468629   .1116816   -0.42   0.676
      polity | -.0222205   .1338221   -0.17   0.869
    lareapc | -.5637075   1.232651   -0.46   0.649
      _cons | -12.55295   20.86813   -0.60   0.549
-----+-----
Instrumented:  inc incsq pwtopen
Instruments:  intery polity lareapc elhsfs incf incfsq
-----+-----
. test inc incsq
( 1) inc = 0      ( 2) incsq = 0
      F( 2,      85) =      0.85   Prob > F =      0.4313

. * Income Peak      .00286067

```

```

. *
. * Energy Depletion
. *
. reg enrдам inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors
Number of obs = 98
F( 6, 91) = 3.22
Prob > F = 0.0065
R-squared = 0.1679
Root MSE = 6.8478

```

enrдам	Coef.	Robust Std. Err.	t	P> t
inc	38.16656	9.120849	4.18	0.000
incsq	-2.206049	.5253217	-4.20	0.000
pwtopen	.024658	.105582	0.23	0.816
intery	-.0044884	.0119176	-0.38	0.707
polity	-.4570515	.1708018	-2.68	0.009
lareapc	.2441835	.4866501	0.50	0.617
_cons	-157.3793	38.71899	-4.06	0.000

```

. test inc incsq ( 1) inc = 0 ( 2) incsq = 0
F( 2, 91) = 8.82 Prob > F = 0.0003

```

```

. * Income Peak 5712.609

```

```

ivreg enrдам (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

```

```

IV (2SLS) regression with robust standard errors
Number of obs = 93
F( 6, 86) = 0.94
Prob > F = 0.4741
R-squared = .
Root MSE = 13.47

```

enrдам	Coef.	Robust Std. Err.	t	P> t
inc	11.48796	53.26458	0.22	0.830
incsq	.6232899	4.809828	0.13	0.897
pwtopen	2.69875	3.787932	0.71	0.478
intery	-.3082474	.4290995	-0.72	0.474
polity	-.8342795	.5914198	-1.41	0.162
lareapc	-2.79859	5.125422	-0.55	0.586
_cons	-131.4493	124.399	-1.06	0.294

```

Instrumented: inc incsq pwtopen
Instruments: intery polity lareapc elhsfs incf incfsq
. test inc incsq ( 1) inc = 0 ( 2) incsq = 0
F( 2, 86) = 2.30 Prob > F = 0.1063

```

```

. * Income Peak .00009948

```

```

. *
. * NO2
. *
. reg no2m inc incsq pwtopen intery polity lareapc, robust

```

```

Regression with robust standard errors
Number of obs = 36
F( 6, 29) = 2.63
Prob > F = 0.0370
R-squared = 0.1574
Root MSE = 41.111

```

no2m	Coef.	Robust Std. Err.	t	P> t
------	-------	------------------	---	------

```

      inc | 422.3331 171.3873 2.46 0.020
      incsq | -23.73463 10.48264 -2.26 0.031
      pwtopen | -.7230931 5.321866 -0.14 0.893
      intery | .0459451 .5636627 0.08 0.936
      polity | -3.200483 1.474742 -2.17 0.038
      lareapc | -5.929677 6.017269 -0.99 0.333
      _cons | -1746.937 670.61 -2.60 0.014
-----
. test inc incsq ( 1) inc = 0 ( 2) incsq = 0

      F( 2, 29) = 4.82 Prob > F = 0.0155

. * Income Peak 7309.8672

ivreg no2m (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

IV (2SLS) regression with robust standard errors      Number of obs = 35
                                                       F( 6, 28) = 7.61
                                                       Prob > F = 0.0001
                                                       R-squared = 0.2287
                                                       Root MSE = 38.595
-----
      no2m |      Coef.   Robust      t   P>|t|
      ----+-----
      inc | 645.1652 237.6681 2.71 0.011
      incsq | -37.44308 14.49659 -2.58 0.015
      pwtopen | -6.302414 6.453029 -0.98 0.337
      intery | .6321251 .6943821 0.91 0.370
      polity | -4.348285 1.010762 -4.30 0.000
      lareapc | -7.102611 6.072446 -1.17 0.252
      _cons | -2611.568 951.1994 -2.75 0.010
-----
Instrumented: inc incsq pwtopen
Instruments: intery polity lareapc elhsfs incf incfsq
-----
. test inc incsq ( 1) inc = 0 ( 2) incsq = 0

      F( 2, 28) = 4.22 Prob > F = 0.0250

. * Income Peak 5515.2865

. *
. * Sulfur Dioxide
. *
. reg sulfdm inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors      Number of obs = 41
                                                       F( 6, 34) = 7.97
                                                       Prob > F = 0.0000
                                                       R-squared = 0.7103
                                                       Root MSE = 22.181
-----
      sulfdm |      Coef.   Robust      t   P>|t|
      ----+-----
      inc | 372.9109 142.1192 2.62 0.013
      incsq | -22.60134 8.396773 -2.69 0.011
      pwtopen | -3.559056 1.653811 -2.15 0.039
      intery | .3531279 .1753988 2.01 0.052
      polity | -6.301547 1.927016 -3.27 0.002
      lareapc | -1.813289 1.711337 -1.06 0.297
      _cons | -1405.922 581.1829 -2.42 0.021
-----
. test inc incsq ( 1) inc = 0 ( 2) incsq = 0

      F( 2, 34) = 4.12 Prob > F = 0.0249

. * Income Peak 3826.6707

ivreg sulfdm (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

IV (2SLS) regression with robust standard errors      Number of obs = 40

```



F( 6, 33) = 7.04  
 Prob > F = 0.0001  
 R-squared = 0.6920  
 Root MSE = 23.207

	Coef.	Robust Std. Err.	t	P> t
inc	513.4604	210.9468	2.43	0.021
incsq	-31.61047	12.99536	-2.43	0.021
pwtopen	-7.005777	4.380388	-1.60	0.119
intery	.7217778	.4694786	1.54	0.134
polity	-6.518679	1.92014	-3.39	0.002
lareapc	-1.443263	2.194518	-0.66	0.515
_cons	-1929.897	829.4044	-2.33	0.026

Instrumented: inc incsq pwtopen  
 Instruments: intery polity lareapc elhsfs incf incfsq

. test inc incsq ( 1) inc = 0 ( 2) incsq = 0  
 F( 2, 33) = 2.97 Prob > F = 0.0654

. \* Income Peak 3366.6849

. \*  
 . \* Suspended Particles  
 . \*  
 . reg suspm inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors  
 Number of obs = 38  
 F( 6, 31) = 11.10  
 Prob > F = 0.0000  
 R-squared = 0.6582  
 Root MSE = 57.823

	Coef.	Robust Std. Err.	t	P> t
inc	645.4215	307.3769	2.10	0.044
incsq	-42.93882	18.05398	-2.38	0.024
pwtopen	-8.707592	5.128309	-1.70	0.100
intery	.9440587	.5607545	1.68	0.102
polity	-7.081393	3.042012	-2.33	0.027
lareapc	-10.76527	5.808484	-1.85	0.073
_cons	-2124.038	1301.286	-1.63	0.113

. test inc incsq ( 1) inc = 0 ( 2) incsq = 0  
 F( 2, 31) = 5.76 Prob > F = 0.0075

. \* Income Peak 1836.4596

ivreg suspm (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

IV (2SLS) regression with robust standard errors  
 Number of obs = 37  
 F( 6, 30) = 9.48  
 Prob > F = 0.0000  
 R-squared = 0.6767  
 Root MSE = 56.411

	Coef.	Robust Std. Err.	t	P> t
inc	729.4699	354.8551	2.06	0.049
incsq	-46.52108	20.75041	-2.24	0.033
pwtopen	-6.292212	4.968944	-1.27	0.215
intery	.643468	.5587868	1.15	0.259
polity	-8.003363	3.546252	-2.26	0.031
lareapc	-12.6961	6.64475	-1.91	0.066
_cons	-2559.111	1497.714	-1.71	0.098

```

Instrumented:  inc incsq pwtopen
Instruments:  intery polity lareapc elhsfs incf incfsq
-----
. test inc incsq      ( 1)  inc = 0      ( 2)  incsq = 0
      F( 2,      30) =    4.41          Prob > F =    0.0209

. * Income Peak      2540.7345

. *
. * Rural access to clean water
. *
. reg ruralan inc incsq pwtopen intery polity lareapc, robust

Regression with robust standard errors                                Number of obs =    57
                                                                    F( 6,      50) =   24.82
                                                                    Prob > F      =   0.0000
                                                                    R-squared     =   0.5979
                                                                    Root MSE     =   18.331
-----
      ruralan |          Coef.   Robust Std. Err.      t    P>|t|
-----+-----
      inc |   -80.11559   37.75428   -2.12   0.039
      incsq |    6.037585   2.284809    2.64   0.011
      pwtopen |   .2180824   .6526446    0.33   0.740
      intery |  -0.0138015   .0830803   -0.17   0.869
      polity |  -0.319822   .5417517   -0.59   0.558
      lareapc |  -9.560386   2.22958   -4.29   0.000
      _cons |   330.8272   156.158    2.12   0.039
-----
. test inc incsq      ( 1)  inc = 0      ( 2)  incsq = 0
      F( 2,      50) =    9.31          Prob > F =    0.0004

. * Income Peak      761.08012

. ivreg ruralan (inc incsq pwtopen = elhsfs incf incfsq) intery polity lareapc, robust

IV (2SLS) regression with robust standard errors                    Number of obs =    55
                                                                    F( 6,      48) =    0.17
                                                                    Prob > F      =   0.9835
                                                                    R-squared     =    .
                                                                    Root MSE     =   128.67
-----
      ruralan |          Coef.   Robust Std. Err.      t    P>|t|
-----+-----
      inc |  -636.2494   6635.896   -0.10   0.924
      incsq |   60.59561   649.4644    0.09   0.926
      pwtopen |    47.396   560.3881    0.08   0.933
      intery |  -5.878212   69.57319   -0.08   0.933
      polity |   .5363244   10.07493    0.05   0.958
      lareapc |  -68.44781   706.2894   -0.10   0.923
      _cons |   1409.271   12966.68    0.11   0.914
-----
Instrumented:  inc incsq pwtopen
Instruments:  intery polity lareapc elhsfs incf incfsq
-----
. test inc incsq      ( 1)  inc = 0      ( 2)  incsq = 0
      F( 2,      48) =    0.02          Prob > F =   0.9849

. * Income Peak      190.55917

* * * * *

```

**Results for interaction of capital/labor and trade (with income in quadratic form).  
cited in R.Ec.Stat. paper**

```

*
. * Benchmark as of 8-12-02. OLS then IV.
. *
. * CO2
. *
. reg co2perc inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors                                Number of obs =    55
                                                                    F( 6, 48) =    35.18
                                                                    Prob > F      =    0.0000
                                                                    R-squared    =    0.7753
                                                                    Root MSE    =    2.6282

-----+-----
      co2perc |           Coef.   Robust          t    P>|t|
              |           Std. Err.          Std. Err.
-----+-----
      inc | -14.71954   7.755569   -1.90   0.064
      incsq |  1.071597   .5011547    2.14   0.038
      pwtopen | -.004053   .0195242   -0.21   0.836
      interkl |  1.08e-06   8.05e-07    1.34   0.187
      polity | -.0095203   .0316964   -0.30   0.765
      lareapc |  .389474   .2660336    1.46   0.150
      _cons |  49.80053  29.69287    1.68   0.100

-----+-----
.. test inc incsq
( 1)  inc = 0
( 2)  incsq = 0

      F( 2, 48) =    10.83                Prob > F =    0.0001

. * Income Peak    961.06426

. ivreg co2perc (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust

IV (2SLS) regression with robust standard errors                    Number of obs =    54
                                                                    F( 6, 47) =    32.59
                                                                    Prob > F      =    0.0000
                                                                    R-squared    =    0.7898
                                                                    Root MSE    =    2.1847

-----+-----
      co2perc |           Coef.   Robust          t    P>|t|
              |           Std. Err.          Std. Err.
-----+-----
      inc | -23.79231   8.417545   -2.83   0.007
      incsq |  1.72546   .5370915    3.21   0.002
      pwtopen | .0102742   .029771    0.35   0.732
      interkl | -8.85e-07   8.31e-07   -1.07   0.292
      polity | -.0272206   .0535811   -0.51   0.614
      lareapc | .3155784   .2587464    1.22   0.229
      _cons |  80.65656  32.46891    2.48   0.017

-----+-----
Instrumented:  inc incsq pwtopen
Instruments:   interkl polity lareapc elhsfs incf incfsq

-----+-----
. test inc incsq
( 1)  inc = 0
( 2)  incsq = 0
      F( 2, 47) =    15.39                Prob > F =    0.0000

. * Income Peak    986.81776

. *
. * Deforestation
. *
. reg defp inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors                                Number of obs =    53
                                                                    F( 6, 46) =     6.41
                                                                    Prob > F      =    0.0001

```

R-squared = 0.3085  
 Root MSE = 1.3138

defp	Coef.	Robust Std. Err.	t	P> t
inc	2.075818	3.032319	0.68	0.497
incsq	-.1691247	.1790041	-0.94	0.350
pwtopen	.011594	.0166507	0.70	0.490
interkl	-1.58e-07	3.11e-07	-0.51	0.614
polity	.0262807	.0555139	0.47	0.638
lareapc	-.0103811	.1037993	-0.10	0.921
_cons	-5.331799	12.51526	-0.43	0.672

```
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
F( 2, 46) = 4.30 Prob > F = 0.0194
```

```
. * Income Peak 462.63861
```

```
. ivreg defp (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust
```

```
IV (2SLS) regression with robust standard errors
Number of obs = 53
F( 6, 46) = 6.56
Prob > F = 0.0000
R-squared = 0.1897
Root MSE = 1.4222
```

defp	Coef.	Robust Std. Err.	t	P> t
inc	4.86595	3.105405	1.57	0.124
incsq	-.3501681	.185161	-1.89	0.065
pwtopen	-.0137896	.0172789	-0.80	0.429
interkl	1.94e-07	3.17e-07	0.61	0.543
polity	.0161798	.0485191	0.33	0.740
lareapc	-.1292735	.1589866	-0.81	0.420
_cons	-14.2882	12.84746	-1.11	0.272

```
Instrumented: inc incsq pwtopen
Instruments: interkl polity lareapc elhsfs incf incfsq
```

```
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
F( 2, 46) = 4.43 Prob > F = 0.0173
```

```
. * Income Peak 1041.0872
```

```
. * Energy Depletion
```

```
. reg enrdam inc incsq pwtopen interkl polity lareapc, robust
```

```
Regression with robust standard errors
Number of obs = 54
F( 6, 47) = 1.48
Prob > F = 0.2046
R-squared = 0.1137
Root MSE = 8.2247
```

enrdam	Coef.	Robust Std. Err.	t	P> t
inc	44.56219	25.38554	1.76	0.086
incsq	-2.683721	1.491049	-1.80	0.078
pwtopen	-.0308318	.021687	-1.42	0.162
interkl	1.09e-06	7.12e-07	1.53	0.132
polity	-.4912701	.3470155	-1.42	0.163

```

      lareapc | .2660432 .6980511 0.38 0.705
      _cons | -176.6527 106.6154 -1.66 0.104
-----+-----
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
      F( 2, 47) = 1.93          Prob > F = 0.1560

. * Income Peak 4033.1887

. ivreg enrdam (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust
IV (2SLS) regression with robust standard errors      Number of obs = 54
                                                    F( 6, 47) = 1.40
                                                    Prob > F = 0.2346
                                                    R-squared = .
                                                    Root MSE = 8.842
-----+-----
      |
      enrdam |      Coef.   Robust
            |           Std. Err.   t   P>|t|
-----+-----
      inc | 71.4296   25.8882   2.76 0.008
      incsq | -4.415733 1.612883  -2.74 0.009
      pwtopen | -.1803119 .1230837  -1.46 0.150
      interkl | 3.92e-06 2.73e-06   1.44 0.157
      polity | -.5708394 .337686  -1.69 0.098
      lareapc | -.4146897 1.042674  -0.40 0.693
      _cons | -270.0651 99.52013  -2.71 0.009
-----+-----
Instrumented: inc incsq pwtopen
Instruments: interkl polity lareapc elhsfs incf incfsq
-----+-----
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
      F( 2, 47) = 3.81          Prob > F = 0.0293

. * Income Peak 3255.43

. *
. * NO2
. *
. reg no2m inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors      Number of obs = 30
                                                    F( 6, 23) = 1.94
                                                    Prob > F = 0.1173
                                                    R-squared = 0.1879
                                                    Root MSE = 43.042
-----+-----
      |
      no2m |      Coef.   Robust
            |           Std. Err.   t   P>|t|
-----+-----
      inc | 410.8601 187.9997   2.19 0.039
      incsq | -22.73852 10.73772  -2.12 0.045
      pwtopen | -.5173574 .4053085  -1.28 0.215
      interkl | 4.71e-07 6.58e-06   0.07 0.944
      polity | -2.728899 3.773059  -0.72 0.477
      lareapc | -7.096475 6.678162  -1.06 0.299
      _cons | -1717.05 808.8863  -2.12 0.045
-----+-----
. test inc incsq
( 1) inc = 0
( 2) incsq = 0
      F( 2, 23) = 2.73          Prob > F = 0.0861

. * Income Peak 8387.1107

ivreg no2m (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust
IV (2SLS) regression with robust standard errors      Number of obs = 30

```

F( 6, 23) = 1.20  
 Prob > F = 0.3406  
 R-squared = 0.1332  
 Root MSE = 44.468

	Coef.	Robust Std. Err.	t	P> t
no2m				
inc	391.1198	381.1072	1.03	0.315
incsq	-20.97587	21.99281	-0.95	0.350
pwtopen	.0783804	.6106091	0.13	0.899
interkl	-.000011	.0000114	-0.97	0.344
polity	-2.868655	4.96476	-0.58	0.569
lareapc	-6.755789	6.347687	-1.06	0.298
_cons	-1698.544	1634.537	-1.04	0.310

Instrumented: inc incsq pwtopen  
 Instruments: interkl polity lareapc elhsfs incf incfsq

. test inc incsq  
 ( 1) inc = 0  
 ( 2) incsq = 0 F( 2, 23) = 1.87 Prob > F = 0.1775

. \* Income Peak 11193.506

. \*  
 . \* Sulfur Dioxide

. reg sulfdm inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors  
 Number of obs = 32  
 F( 6, 25) = 3.57  
 Prob > F = 0.0108  
 R-squared = 0.5139  
 Root MSE = 15.76

	Coef.	Robust Std. Err.	t	P> t
sulfdm				
inc	303.3421	147.0133	2.06	0.050
incsq	-18.11208	8.846506	-2.05	0.051
pwtopen	-.3827739	.2273838	-1.68	0.105
interkl	3.92e-06	4.04e-06	0.97	0.342
polity	-1.730975	2.339441	-0.74	0.466
lareapc	-3.323324	1.480258	-2.25	0.034
_cons	-1188.958	607.9216	-1.96	0.062

. test inc incsq  
 ( 1) inc = 0  
 ( 2) incsq = 0 F( 2, 25) = 2.26 Prob > F = 0.1250

. \* Income Peak 4333.0456

. ivreg sulfdm (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust

IV (2SLS) regression with robust standard errors  
 Number of obs = 32  
 F( 6, 25) = 3.66  
 Prob > F = 0.0095  
 R-squared = 0.5088  
 Root MSE = 15.842

	Coef.	Robust Std. Err.	t	P> t
sulfdm				
inc	334.4655	150.8396	2.22	0.036
incsq	-19.95198	9.01319	-2.21	0.036
pwtopen	-.3429395	.2848613	-1.20	0.240
interkl	3.54e-06	4.79e-06	0.74	0.467
polity	-1.486867	2.394407	-0.62	0.540

```

      lareapc | -3.205778   1.549381   -2.07   0.049
      _cons | -1323.283   627.9271   -2.11   0.045
-----+-----
Instrumented:  inc incsq pwtopen
Instruments:  interkl polity lareapc elhsfs incf incfsq
-----+-----
. test inc incsq
( 1)  inc = 0
( 2)  incsq = 0
      F( 2, 25) = 2.46          Prob > F = 0.1059

. * Income Peak 4366.7006

. *
. * Suspended Particles
. *
. reg suspm inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors          Number of obs = 30
                                                F( 6, 23) = 4
                                                Prob > F = 0.0067
                                                R-squared = 0.5523
                                                Root MSE = 56.432
-----+-----
      |               Robust
suspm |               Coef.  Std. Err.      t    P>|t|
-----+-----
      inc | 551.3948   508.8648     1.08   0.290
      incsq | -35.79599  29.24197    -1.22   0.233
      pwtopen | -.6467419  .7005434    -0.92   0.365
      interkl | 9.64e-06   .0000133     0.73   0.475
      polity | -2.355989  7.672784    -0.31   0.762
      lareapc | -10.04963  5.408642    -1.86   0.076
      _cons | -1889.7    2157.393    -0.88   0.390
-----+-----
. test inc incsq      ( 1)  inc = 0          ( 2)  incsq = 0
      F( 2, 23) = 5.39          Prob > F = 0.0120

. * Income Peak 2212.5653

ivreg suspm (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust

IV (2SLS) regression with robust standard errors          Number of obs = 30
                                                F( 6, 23) = 3.93
                                                Prob > F = 0.0076
                                                R-squared = 0.5400
                                                Root MSE = 57.199
-----+-----
      |               Robust
suspm |               Coef.  Std. Err.      t    P>|t|
-----+-----
      inc | 673.5513   612.294     1.10   0.283
      incsq | -42.70518  35.24489    -1.21   0.238
      pwtopen | -.2839184  .9247932    -0.31   0.762
      interkl | 4.46e-06   .0000192     0.23   0.818
      polity | -2.768539  8.20899     -0.34   0.739
      lareapc | -9.277383  5.315844    -1.75   0.094
      _cons | -2436.355  2587.462    -0.94   0.356
-----+-----
Instrumented:  inc incsq pwtopen
Instruments:  interkl polity lareapc elhsfs incf incfsq
-----+-----
. test inc incsq
( 1)  inc = 0
( 2)  incsq = 0
      F( 2, 23) = 2.95          Prob > F = 0.0722

. * Income Peak 2659.9431

```

```

. *
. * Rural access to clean water
. *
. reg ruralan inc incsq pwtopen interkl polity lareapc, robust

Regression with robust standard errors
Number of obs = 32
F( 6, 25) = 10.68
Prob > F = 0.0000
R-squared = 0.6315
Root MSE = 20.134

```

ruralan	Coef.	Robust Std. Err.	t	P> t
inc	-33.88171	62.67165	-0.54	0.594
incsq	2.889868	3.786618	0.76	0.452
pwtopen	.0391187	.214443	0.18	0.857
interkl	3.71e-06	5.01e-06	0.74	0.466
polity	.1281525	.9003763	0.14	0.888
lareapc	-9.772383	3.517579	-2.78	0.010
_cons	165.3194	255.6344	0.65	0.524

```

. test inc incsq ( 1) inc = 0 ( 2) incsq = 0
F( 2, 25) = 3.61 Prob > F = 0.0420

```

```

. * Income Peak 351.48102

```

```

. ivreg ruralan (inc incsq pwtopen = elhsfs incf incfsq) interkl polity lareapc, robust

```

```

IV (2SLS) regression with robust standard errors
Number of obs = 32
F( 6, 25) = 9.87 Prob > F = 0.0000
R-squared = 0.5625
Root MSE = 21.937

```

ruralan	Coef.	Robust Std. Err.	t	P> t
inc	30.25503	78.07504	0.39	0.702
incsq	-1.3484	5.00639	-0.27	0.790
pwtopen	-.3831083	.4671899	-0.82	0.420
interkl	.0000108	9.99e-06	1.08	0.289
polity	-.1202621	.9388455	-0.13	0.899
lareapc	-11.92875	4.403313	-2.71	0.012
_cons	-46.92236	283.4336	-0.17	0.870

```

Instrumented: inc incsq pwtopen
Instruments: interkl polity lareapc elhsfs incf incfsq

```

```

. test inc incsq ( 1) inc = 0 ( 2) incsq = 0
F( 2, 25) = 1.00 Prob > F = 0.3810

```

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. * Income Peak 74523.099

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* * * * *

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## Endnotes

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<sup>1</sup> References in the wider debate on globalization and the environment are given in Frankel (2003).

<sup>2</sup> Appendix Figure A1 shows the graph for all three measures of air pollution used in this study.

<sup>3</sup> One way to see the ambiguity is to distinguish three channels whereby trade, like any other determinant of real income, can affect the environment. The *scale* effect is the obvious channel whereby higher GDP leads to higher pollution. But in the language of, e.g., Copeland and Taylor (2003) and Grossman and Krueger (1993), there is also a *composition* effect (e.g., agriculture vs. manufacturing vs. services have different effects on the environment) and a *technique* effect (any given sector can use cleaner or dirtier techniques of production). The question is whether the latter two effects can outweigh the first. The literature is surveyed in Dean (1992, 2001).

<sup>4</sup> It is important to emphasize a key difference between the race to the bottom hypothesis and the pollution haven hypothesis: while the former implies a negative effect on the overall world level of environmental regulation, the latter does not. Some countries may choose high environmental standards for their own production, and import from others goods that embody pollution. The second group can be said to exploit or develop a comparative advantage in pollution. The pollution haven hypothesis with respect to trade is tested toward the end of this paper. Jaffe, Peterson, Portney and Stavins (1995) survey the literature with regard to direct investment in the U.S..

<sup>5</sup> References for these three interrelated hypotheses include Braithwaite and Drahos (2000), Eskeland and Harrison (2002), Esty and Gentry (1997), Schmidheiny (1992), and Vogel (1995).

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<sup>6</sup> The environmental gains from trade, even for a given level of GDP per capita, would occur particularly if measured GDP does not adequately capture the increase in welfare arising from enhanced variety of consumption.

<sup>7</sup> Notice in Figure 1 that the low-democracy countries tend to have higher SO<sub>2</sub> pollution. Barrett and Graddy (2000) also find that an increase in civil and political freedoms significantly reduces some measures of pollution.

<sup>8</sup> They use the Dollar index of trade distortion, a measure of relative prices, which has been heavily criticized by Rodriguez and Rodrik (2001).

<sup>9</sup> Rodriguez and Rodrik (2001) also criticize her measure of trade distortion, the black market premium.

<sup>10</sup> Levinson (1999) shows that controlling for endogeneity of environmental regulation can change results, in his study of hazardous waste trade. Dean (2002) treats income as endogenous in her study of the effect of trade liberalization on water pollution across Chinese provinces.

<sup>11</sup> While a number of studies have confirmed the EKC, especially for SO<sub>2</sub> and Particulate Matter, the results are not always favorable; e.g., Bradford, Schlieckert and Shore (2000) get mixed answers. Many more EKC references are available there, in Frankel (2002), or in the working paper version of the present study.

<sup>12</sup> Rodrik (1995) and Rodriguez and Rodrik (2001) are among those critical of previous studies on the grounds of simultaneity.

<sup>13</sup> The usual presumption is that environmental regulation, by raising business costs, slows economic growth. But we should also consider the Porter Hypothesis, according to which a tightening of environmental regulation stimulates technological innovation and thereby raises productivity. (E.g., Porter and van der Linde, 1995.)

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<sup>14</sup> The usual presumption is that environmental regulation, by raising business costs, slows economic growth. On the other hand the Porter Hypothesis (in which a tightening of environmental regulation stimulates technological innovation and thereby raises productivity) has the opposite implication. See e.g., Porter and van der Linde, 1995.

<sup>15</sup> Energy depletion is a measure computed for the World Bank's *World Development Indicators*. It is equal to the product of unit resource rents and the physical quantities of fossil fuel energy extracted (including coal, crude oil, and natural gas). Table 3.15, available at [http://www.worldbank.org/data/wdi2001/pdfs/tab3\\_15.pdf](http://www.worldbank.org/data/wdi2001/pdfs/tab3_15.pdf), explains the data computations.

<sup>16</sup> That is, it is estimated that increases in income in the low-income countries increase pollution, and in the high-income countries reduce it. The coefficient on openness is again negative for all three measures of air pollution.

<sup>17</sup> In most cases, the effects of polity, area, and quadratic income -- not reported here, to save space -- go in the same direction as with the air pollution indicators; the EKC shows up highly significant for deforestation, energy depletion, and rural water access.

<sup>18</sup> Further, the coefficient on quadratic income is positive and highly significant, while in the spline version income has a positive effect through all three segments in this case. This confirms others' findings of no EKC for CO<sub>2</sub>.

<sup>19</sup> The significance level for SO<sub>2</sub> is 5% under OLS and 10% under IV (and for PM is more marginal; all results available in a working paper). This is consistent with the finding of Antweiler, Copeland and Taylor (2001) that trade has a significantly less favorable effect on SO<sub>2</sub> emissions in rich countries than in poor countries. Their explanation is that, because rich countries have higher capital/labor ratios, the factor-based pollution-haven effect -- the third hypothesis, considered below -- tends to outweigh the income-based pollution-haven effect.

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<sup>20</sup> E.g., Copeland and Taylor (2003) and Grossman and Krueger (1993, 1995), who suggest that the overall effect might be a reduction of pollution in the aggregate, contrary to the popular view of “eco-dumping” in poor countries.

<sup>21</sup> E.g., Frankel and Romer (1999), or Irwin and Tervio (2002), or the working paper version of this paper.